Multidecadal climate variability over northern France during the past 500 years and its relation to large-scale atmospheric circulation

Dieppois, B, Lawler, D, Slonosky, V, Massei, N, Bigot, S, Fournier, M & Durand, A

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Complete List of Authors:	Dieppois, Bastien; University of Cape Town, African Climate Development and Initiative Lawler, Damian; Coventry University, Centre for Agroecology, Water and Resilience (CAWR) Slonosky, Victoria; anadian Historical Climate Data Project, Data Analysis Massei, N; UMR 6143 M2C, University of Rouen, Department of Geology; Bigot, Sylvain; Université de Grenoble, Geography; Fournier, Matthieu; UMR CNRS 6143 Morhodynamique Continentale et Côtière, Université de Rouen, FED 4116 SCALE, Geology Durand, Alain; CNRS UMR 6143, Geology
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Multidecadal climate variability over northern France during the past 500 1 years and its relation to large-scale atmospheric circulation 2 3 Dieppois B.^{1, 2}, Lawler D.M.¹, Slonosky V.³, Massei N.², Bigot S.⁴, Fournier M.², Durand A.² 4 bastien.dieppois@univ-rouen.fr 5 6 ¹ Centre for Agroecology, Water and Resilience (CAWR), Coventry University, Coventry 7 CV1 5FB, United Kingdom 8 ² Laboratoire Morphodynamique Continentale et Côtière (M2C), Université de Rouen, CNRS 9 UMR 6143, FED 4116 SCALE, Mont-Saint Aignan, France 10 ³ Canadian Historical Data Rescue, Montreal, Canada 11 ⁴ Laboratoire d'étude des Transferts en Hydrologie et Environnement (LTHE), Université 12 Grenoble Alpes (UGA), Grenoble, France 13 14 Running title: Multidecadal climate variability over northern France 15 16 17 Abstract: We examine secular changes and multidecadal climate variability on a seasonal 18 19 scale in northern France over the last 500 years, and examine the extent to which they are driven by large-scale atmospheric variability. Multiscale trend analysis and segmentation 20 procedures show statistically significant increases of winter and spring precipitation amounts 21 in Paris since the end of the 19th century. This changes the seasonal precipitation distribution 22 from one with a pronounced summer peak at the end of the Little Ice Age to an almost 23 uniform distribution in the 20th century. This switch is linked to an early warming trend in 24 winter temperature. Changes in spring precipitation are also correlated with winter 25 precipitation for time-scales greater than 50 years, which suggests a seasonal persistence. 26 27 Hydrological modelling results show similar rising trends in river flow for the Seine at Paris. 28 However, such secular trends in the seasonal climatic conditions over northern France are substantially modulated by irregular multidecadal (50-80 year) fluctuations. Furthermore, 29 since the end of the 19th century, we find an increasing variance in multidecadal hydroclimatic 30 31 winter and spring, and this coincides with an increase in the multidecadal North Atlantic 32 Oscillation (NAO) variability, suggesting a significant influence of large-scale atmospheric 33 circulation patterns. However, multidecadal NAO variability has decreased in summer. Using Empirical Orthogonal Function analysis, we detect multidecadal North Atlantic sea-level 34

pressure anomalies, which are significantly linked to the NAO during the Modern period. In particular, a south-eastward (south-westward) shift of the Icelandic Low (Azores High) drives substantial multidecadal changes in spring. Wetter springs are likely to be driven by potential changes in moisture advection from the Atlantic, in response to northward shifts of North Atlantic storm tracks over European regions, linked to periods of positive NAO. Similar, but smaller, changes in rainfall are observed in winter.

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Keywords: Northern France, Precipitation, Climate, River flow changes, Multidecadal
variability, North Atlantic Oscillation, Atmospheric circulation, Non-stationarity

44

45 **1. Introduction**

While anthropogenic influence may contribute to recent and future temperature and 46 precipitation trends, climate changes prior to the Industrial Revolution in the 18th century can 47 be attributed to natural causes such as changes in solar activity, volcanic eruptions and natural 48 changes in greenhouse gas concentration (e.g., Jansen et al., 2007). In this context, knowledge 49 50 concerning the range of variability in hydroclimatic variables such as seasonal precipitation, 51 drought and river flow in past centuries at regional scales is important (e.g., Lawler, 1987; Slonosky, 2002; Xoplaki et al., 2005; Pauling et al., 2006). Furthermore, linking such changes 52 53 to large-scale atmospheric and oceanic oscillations is crucial to building an understanding of 54 their controls.

55

European climate is known to be strongly related to the state of the atmospheric circulation
over the North Atlantic (*e.g.*, Hurrell, 1995; Slonosky *et al.*, 2001; Cassou *et al.*, 2004;
Xoplaki *et al.*, 2005; Hurrell and Deser, 2009; Küttel *et al.*, 2010). More specifically, the
North Atlantic Oscillation (NAO) is described as a key driver of climate conditions over

Europe (Hurrell, 1995; Cassou et al., 2004; Hurrell and Deser, 2009), and is likely to 60 61 influence hydrological systems (Philips et al., 2003; Kingston et al., 2006a, b, 2009; Massei et al., 2007; Lavers et al., 2010; Holman et al., 2011). The NAO can be described as an 62 oscillation of atmospheric mass between the northern North Atlantic and the subtropical 63 Atlantic, usually defined through changes in surface pressure (Hurrell, 1995; Cassou et al., 64 2004; Hurrell and Deser, 2009). This oscillation, which is the only teleconnection pattern 65 evident throughout the year in the northern Hemisphere (Barnston and Livezey, 1987; Hurrell 66 67 et al., 2003), produces changes in wind speed and direction affecting heat and moisture transport between the Atlantic and the neighbouring European continent (Hurrell, 1995; 68 69 Cassou et al., 2004; Hurrell and Deser, 2009).

70

However, the links between NAO pattern and European hydroclimate are complex and poorly 71 72 understood. In particular, the relationship are not constant over time, and show seasonal as well as long-term temporal nonstationarity (Slonosky and Yiou, 2002; Raible et al., 2006; 73 Gamiz-Fortis et al., 2008; Fritier et al., 2012; Lehner et al., 2012). For instance, while 74 positive winter NAO drives wetter conditions over northwestern Europe (Hurrell, 1995), 75 76 positive summer NAO, which is displaced significantly north-eastward compared to winter, implies dryer conditions (Hurrell and Deser, 2009; Folland et al., 2009). Similar changes in 77 the position of the poles of the NAO also occur over longer time-scales, but the reasons are 78 79 less well understood. According to Raible et al. (2014), who examined several control and 80 transient millennium-scale simulations with coupled models, shifts in the centres of action of 81 the NAO are not related to changes in external forcing. This is consistent with internal climate variability, such as North Atlantic sea-ice-atmosphere interactions (Hilmer and Jung, 2000; 82 83 Jung et al., 2003). However, this contradicts other studies suggesting shifts in the NAO pattern under greenhouse-gas-induced warming (Ulbrich and Christoph, 1999; Dong et al., 84

2011), or in response to fluctuations in solar activity (Ineson *et al.*, 2011; van Loon *et al.*,
2012).

87

In particular, European climate between the 16th and 19th centuries has seen cold multi-88 decadal periods, within the period known as the "Little Ice Age" (LIA; Grove, 1988; Bradley 89 90 and Jones, 1993; Jones et al., 1998; Luterbacher et al., 2004; Büntgen and Hellman, 2014). 91 The most popular hypothesis of LIA climate change is a weakening of the Atlantic Meridional 92 Overturning Circulation (AMOC) in response to a pervasive negative NAO, which implies both dry and cold anomalies over the northern regions of Europe (Luterbacher et al., 1999; 93 Trouet et al., 2009; Mann et al., 2009; Trouet et al., 2012). However, contradictory evidence 94 95 from other proxy-based studies suggests that the LIA was characterized by enhanced storminess over the northern North Atlantic (Meeker and Mayeki, 2002; Dawson et al., 2003, 96 97 2007), which is usually consistent with a positive NAO. Such multidecadal fluctuations in the NAO have also been reported over the last two centuries, hereafter referred to as the Modern 98 period (Cook et al., 1998; Goodkin et al., 2008; Olsen et al., 2012; Dieppois et al., 2013; Sun 99 100 et al., 2015), and in precipitation amounts in northern France (Dieppois et al., 2013) as well as 101 French river flow (Boé and Habets, 2014). Analysing the annual relationship between NAO 102 and precipitation in England and northern France back to the pre-industrial period, Dieppois 103 et al. (2013) show a phase change at the multidecadal time-scale between the LIA and the 104 Modern period: before 1850, negative NAO was associated with greater precipitation but, 105 after 1850, positive NAO was associated with greater precipitation. This non-stationarity in 106 the annual relationship might result from long-term changes in seasonal NAO patterns, but 107 these have not yet been explored. This issue is especially problematic since multidecadal 108 fluctuations in European and northern France hydroclimate, and in particular river flows, may 109 also have serious impacts for society in influencing recent and future trends of water110 resources.

111

In this paper, using long-term climatic observations, reconstructions and conceptual 112 hydrologic modelling back to 1500, we examine: i) how the annual cycles of hydroclimatic 113 114 variables (precipitation, temperature, potential evapotranspiration [PE]) affecting northern 115 France have changed since the end of the LIA; ii) how the multidecadal changes in the 116 relationship between NAO and precipitation over northern France are related to changes in the seasonal relationship; iii) how river flow is affected by such changes, by examining 117 118 streamflow of one of the main French rivers, the Seine River flow; iv) how North Atlantic 119 Sea-Level Pressure (SLP) anomalies affected the multidecadal variability of the NAO after 1850. 120

121

The paper is organised as follows. In Section 2, we discuss the datasets used in this study, 122 before we describe the analysis methods in Section 3. Seasonal changes and trends in 123 124 hydroclimatic variables over the past 500 years are examined in Section 4.1. In Section 4.2, 125 we investigate coherent seasonal changes in the multidecadal variability of northern France hydroclimate and of the NAO. Potential effects of such changes on Seine river flow are 126 127 estimated by hydrological modelling in Section 4.3. The multidecadal North Atlantic SLP 128 anomalies are then extracted for each season in Section 5.1, before describing their potential 129 impacts on European precipitation during the Modern period (1820-2000, hereafter) in 130 Section 5.2. Our main results are interpreted and their wider implications discussed in Section 6. 131

132

133 **2. Data**

134 **2.1. Precipitation data**

135 All datasets used in this paper are summarised schematically in Figure 1, together with their corresponding timescales. The Paris observations of monthly precipitation totals used for the 136 137 study originate from updates to Slonosky (2002). They come from two main sources: the yearly summaries of the weather which were published in the Mémoires de l'Academie 138 139 Royale des Sciences from 1688 to 1754, and previously published data collected by Emilien 140 Renou (1815–1902) from 1806 to 1902 (Renou, 1885). Such time series can be affected by 141 modifications of measurement conditions such as rain gauge displacements, especially height and exposure, measurement instrument replacement, or changes in rain gauge environment, 142 143 which could result in artificial shifts that do not reflect climate variations. However, as suggested by Slonosky (2002) and Dieppois et al. (2013), this time series, once adjusted (e.g., 144 Tabony 1980, 1981), can be considered homogeneous. The spatial coherence of interannual 145 146 Paris precipitation anomalies with larger scale precipitation fields between 1901 and 2009 has nevertheless been estimated empirically for each season in Figure 2 (*i.e.*, winter: DJF; spring: 147 MAM; summer: JJA; autumn: SON). This is based on a pointwise correlation with high-148 TS 3.10.1 monthly 149 resolution grids from the CRU datasets (explained at 150 badc.nerc.ac.uk/view; Fig.1). The year-to-year variations of seasonal Paris precipitation series 151 are significantly correlated with a large area of precipitation observations from the CRU data 152 set over NW Europe with a maximum centred on Paris (Fig. 2a-d), suggesting robust spatial 153 coherence of the data in all four seasons.

154

We also use the high-resolution precipitation reconstruction of Pauling *et al.* (2006) which covers the European land areas $(30^{\circ}N - 70^{\circ}N; 30^{\circ}W - 40^{\circ}E)$ resolved on a $0.5^{\circ} \times 0.5^{\circ}$ grid (Fig.1). This reconstruction is based on transfer functions derived from Empirical Orthogonal Function (EOF) regression between predictors, such as long instrumental precipitation series or natural proxies, and the predictand (CRU TS 3.10 dataset). We established a northern France precipitation index from this reconstructed field using an average over latitudes $47.5^{\circ}N - 50.5^{\circ}N$ and longitudes $0 - 4^{\circ}E$. By comparing the reconstructed precipitation to the Paris observations with seasonal correlation maps, the reconstructed northern France precipitation index emerges as a useful, spatially coherent, representation of year-to-year fluctuations of northwestern European precipitation for all seasons (Fig. 2e-h).

165

Note that precipitation measurements from the Paris Observatory were used as predictor in the
Pauling reconstruction (Pauling *et al.*, 2006) and, thus, are not independent observations.
However, in the present study, the observed Paris precipitation and the reconstructed northern
France index is treated independently to avoid any circular argument.

170

171 **2.2. Temperature and Potential Evapotranspiration data**

172 The homogenized Central England Temperature index, which is the longest instrumental 173 record of temperature in the world, has been used to examine temperature fluctuations that 174 may have influenced the climate of northern France over the past five centuries (Fig.1). This 175 monthly time-series, covering 1659 to 1973, was compiled by Manley (1974), before being 176 updated by Parker et al. (1992). Although this index is representative of a roughly triangular 177 area of the United Kingdom enclosed by Lancashire, London and Bristol, its interdecadal to 178 multidecadal variability is almost identical to that of Paris temperature records (1757 – 179 present; Fig.1), and the trend of the same order (Dieppois et al., 2013). Their seasonal correlations are between 0.7 and 0.82 at $p \le 2.2^{e-16}$, with maximum and minimum in winter 180 181 and autumn, respectively. We note differences in the mean, especially between spring and autumn (Paris T°C minus CET = 0.36°C in winter, -1.90°C in spring, -2.71°C in summer, -182

1.24°C in autumn). Very similar results were, however, obtained using Paris temperature time
series (as discussed hereafter).

185

186 We also calculated the potential evapotranspiration (PE, in mm), which forms a key part of catchment water budgets (Fig.1). PE is defined as the amount of evaporation and transpiration 187 188 that would occur if sufficient water sources were available, *i.e.*, when soil moisture is not a 189 limiting factor. The complexity of these calculations varies greatly, ranging from a simple 190 function of just one atmospheric variable, often temperature (Thornthwaite, 1948), to those requiring a range of variables, such as relative humidity, wind speed and net solar radiation 191 192 (e.g., Hargreaves, 1994; Droogers and Allen, 2002). The *Thornthwaite* equation, which can be 193 used when only temperature data are available, has been selected, and is given by:

194
$$PE = 16N_m \left(\frac{10\overline{T}_m}{I}\right)^a mm$$
 (1)

where m is the index for months with m = 1, 2, 3...12, N_m is the monthly adjustment factor related to hours of daylight, \overline{T}_m is the monthly mean air temperature (C), the exponent *a*, and *I* is the heat index for the year, given by :

198 I =
$$\sum i_m = \sum \left(\frac{\overline{T}_m}{5}\right)^{1.5}$$
 (2)
199 $a = 6.7 \times 10^{-7} \times I^3 - 7.7 \times 10^{-5} \times I^2 + 1.8 \times 10^{-2} \times I + 0.49$ (3)

200

Estimates of PE were very similar using Central England or Paris temperatures (*Corr.[R]* \geq 0.7; -11.6 mm \leq *Mean Error [ME]* \leq 2.3 mm; the ratio of Standard Devations [*rSD*] \geq 0.66). Note, however, that such a simplified model of PE may respond with less accuracy to a warming climate than calculations based on the underlying physical principles, which take into changes relative humidity, wind speed and net solar radiation (Sheffield *et al.*, 2012).

206

207 **2.3. Sea-Level Pressure data**

208 In the atmospheric circulation analysis, we compare two observed and two reconstructed SLP fields over the North-Atlantic-European area (30°N - 70°N; 30°W - 40°E; Fig.1). The 209 HadSLP2r dataset compiled by the Met Office Hadley Centre, which is resolved on a $5^{\circ} \times 5^{\circ}$ 210 211 grid between 1850 and 2013, has been used as a reference over the instrumental period (Fig. 1; Allan and Ansell, 2006). We also used an extended observation field from the Annual to 212 213 Decadal Variability In Climate in Europe (ADVICE) project (Fig. 1), which is based on a 214 compilation and homogenization of 51 European stations with starting dates ranging from 215 1755 to 1871 (Jones et al., 1999). We used reconstructions performed by Luterbacher et al. (2002) and Küttel *et al.* (2010), which are both resolved on a $5^{\circ} \times 5^{\circ}$ grid but cover different 216 217 periods (Fig. 1). These two reconstructions were computed by means of an EOF regression 218 technique with predictors comprising early instrumental time series as well as several climate indices based on documentary proxy data from various sites in Europe. However, the 219 reconstruction from Küttel et al. (2010), which only uses terrestrial instrumental pressure 220 221 series and maritime wind information derived from ship logbook data, is more reliable over the eastern North Atlantic than the other early-instrumental SLP fields (Küttel et al., 2010). It 222 223 should also be noted that precipitation (but not the Paris monthly observations) was used as 224 predictor in the Luterbacher et al. (1999) reconstruction.

225

From these SLP datasets, the principal component (PC) time series of the leading EOF of reconstructed SLP anomalies over the Euro-Atlantic sector are calculated seasonally (*i.e.*, winter: DJF; spring: MAM; summer: JJA; autumn: SON) to define NAO indices. This ensures the ability to track the movement of the NAO centres of action through the annual cycle, and thus to provide an optimal representation of each seasonal NAO spatial pattern. Note that the varying start dates and lengths of the observed and reconstructed datasets lead to slightly different definitions of the Modern period (Fig. 1). 233

3. Methods

3.1. Estimation of seasonal and multidecadal variabilities

236 First, we examined the modifications of the seasonal temperature anomalies and of the seasonal precipitation distribution using a cartographic representation of observed and 237 238 reconstructed values over centuries. Second, changes in the mean value of average seasonal 239 temperature, PE, and precipitation were investigated using the segmentation procedure of 240 Aksoy et al. (2008). This method uses a least squares algorithm to detect optimal breaks between adjacent data-segments with means that are significantly different using the Scheffe 241 242 contrast test at p = 0.05. Third, monotonic trends in each seasonal average were calculated 243 using a modified Mann-Kendall trend test accounting for serial correlation (Hamed and Rao, 1998). To eliminate the effect of serial correlation, the effective sample size, which allows us 244 245 to modify the Mann-Kendall S statistics, has been estimated according to a theoretical relationship based on first-order auto-regressive (AR[1]) model of the raw time series 246 considered. In addition, to examine the influence of potential multidecadal fluctuations on 247 trends (*i.e.*, the Sen's slope values; Sen, 1968), a rolling trend analysis across every possible 248 249 time-scale has been used (Mc Cabe and Wolock, 2002; Liebman et al., 2010). The same approach has been employed to examine whether winter trends and variability in temperature 250 251 and precipitation are statistically related to those of spring and summer months, and also if 252 this is restricted to specific time-scales.

253

Potentially synchronous changes in precipitation, temperature, PE and NAO on multidecadal time-scales are then examined by submitting each seasonal time series to the Continuous Wavelet Transform (CWT). By representing the time series in the time-scale space, one can determine which scales of variability (or periods in a Fourier sense) are the dominant

variability modes through time. It should be noted, however, that multi-proxy climate 258 259 reconstructions using tree-ring proxies as predictors are likely to overestimate low-frequency signals (Franke et al., 2013). Such a decomposition of each seasonal signal is conducted with 260 261 a Morlet mother wavelet with angular frequency 6 to produce the local wavelet spectra, which provides a good trade-off between time and frequency resolution. The detailed explanations of 262 263 CWT methodology and its applications to the analysis of hydrological and climatic signals are now widely documented (e.g., Torrence and Compo 1998; Maraun, 2006; Sang, 2013). The 264 265 significance test of the wavelet spectrum for geophysical signals generally assumes a red noise background spectrum for the null hypothesis. The wavelet spectrum is tested for every 266 267 point in time and scale to check whether the power exceeds a certain critical value determined by Monte-Carlo simulations of AR[1] processes. The cone of influence, which delineates the 268 269 area under which power can be underestimated as a result of edge effects and zero padding, is 270 also calculated and represented on all spectra as a thick bold line.

271

3.2. Long-term Seine river flow reconstruction

273 The potential effects of secular and multidecadal changes in precipitation, temperature and PE 274 on the Seine River flow at Paris have been evaluated by hydrological modelling using GR2M (Génie Rural à 2 paramètres au pas de temps Mensuel; Mouelhi et al., 2006). GR2M is a two-275 276 store empirical lumped hydrological model running, which runs on a monthly time-step basis 277 with two parameters: (a) the maximum capacity of the production store (S); and (b) the water 278 exchange term with neighbouring catchments (*i.e.*, hydrographic network and aquifer) which 279 applies to the routing store (R). The model is forced by monthly precipitation (P) and potential evapotranspiration (*PE*) and returns a monthly flow (Q). Note that the soil moisture 280 281 storage capacity (parameters for store S) controls the response of the model to rain events, and to a certain extent, the variability of the simulated flow. As soil moisture storage capacity 282

increases, the simulated flow depends less on the current rainfall and more on the storage level, itself dependent on past rainfall. To avoid any problems related to this issue our model initialization was defined according to Mouelhi *et al.* (2006). Note that very similar simulated river flow variability was obtained using Central England or Paris temperature ($R \ge 0.98$ and $rSD \ge 0.91$). However, using Central England temperature leads to underestimate simulated flow by a maximum of 22.64 m³.s⁻¹ (~7.5%).

289

290 The simulated Paris river flow is then compared with the observed river flows from a gauging station in south-east Paris at Austerlitz, from 1885 to 2014. Metadata indicate that this river 291 292 flow gauge is currently strongly influenced by human activity, but shows a good quality control (http://www.hydro.eaufrance.fr/). Artificial influence is principally related to 293 reservoirs near the headwaters, built in 1973, which control up to 17% of river inputs. 294 295 However, in our tests, we found no changes in homogeneity of the time series in 1973 (cf. 296 Sect. 4.3). Because of differences between Central England and Paris temperatures (cf. Sect 2.2; contributing to underestimate simulated flow by a maximum of 7.5%), seasonal averages 297 of simulated river flow have been adjusted according to the calculated mean error (ME; 298 winter: -112.5 m³.s⁻¹; spring: 6.3 m³.s⁻¹; summer: 162.8 m³.s⁻¹). The goodness-of-fit between 299 observed and simulated Paris river flows has been evaluated over the common period (1885– 300 301 2009). By examining the ratio of Standard Devations (rSD), the simulated variability is 302 slightly underestimated by a factor of two in winter (rSD: 0.47) and spring (rSD: 0.6). 303 Simulated and observed variability are similar in summer (rSD: 1.05). The timings of simulated and observed fluctuations are very consistent in winter and spring with $R \ge 0.74$, 304 305 and they are reasonably good in summer (R = 0.58). The agreement index, which describes 306 additive and proportional differences in the observed and simulated means and variances, and the Volumetric Efficiency (VE), suggest a good match of simulated to observed Seine river 307

flow. *RSME* and *NSE* coefficients, which are linked to the mean, have been substantially reduced in winter and spring by the mean corrections. The Kling-Gupta Efficiency (*KGE*) coefficients, which facilitate the analysis of the relative importance of its different components (correlation, bias and variability), also ensure that the bias and variability ratios are not cross-correlated.

313

314 **3.3. Multidecadal SLP patterns and their impacts on European precipitation**

315 To improve our understanding of the multidecadal NAO variability, multidecadal time-scale 316 anomalies of North Atlantic SLP have been extracted using EOFs (Preisendorfer, 1988). 317 Across the Modern period (1820 –2000 for reconstructed data or 1850–2013 for observational 318 data), we first subtracted a linear trend from all grid points and then used a 30-year running 319 mean to define a primitive low-pass filter over each grid point. The first-EOFs were then 320 computed for each season and, subsequently, compared to those produced using unfiltered 321 HadSLP2r dataset, *i.e.*, the observed seasonal NAO patterns. The statistical relationships between multidecadal anomalies in the North Atlantic SLP anomalies and in the NAO 322 323 patterns were quantified with spatial and temporal correlations.

324

325 We then examined changes in the seasonal relationships between the multidecadal North 326 Atlantic SLP anomalies (which are potentially linked to multidecadal NAO variability) and 327 European precipitation during the Modern period. Maps of the Pearson's correlation 328 coefficients between the 30-year running means of European precipitation and multidecadal 329 North Atlantic SLP anomalies have been computed for each season. The degrees of freedom 330 (DOF) in the local significance calculations at p = 0.05 were adjusted using the estimated 331 decorrelation scales. To examine whether the map is statistically distinguishable from random noise, we also calculated the field significance according to the binomial procedure developed 332

by Livezey and Chen (1983). Note that the problem of statistical independence between reconstructions of Luterbacher *et al.* (2002) and of Pauling *et al.* (2006) mainly affected the results prior to the Modern period. These two reconstructions used the NCEP/NCAR SLP

data and observed precipitation field, respectively, as predictands over the 20th century.

337

4. Seasonal and multidecadal hydroclimatic variability

339 4.1 Time evolution of seasonal fluctuations

Figure 3a shows the seasonal distribution of CET anomalies since 1659. The coldest decadal 340 period observed at the end of the 17th century is clearly distinct from the recent warmest 341 342 period of the post-1970s. This is a feature of all seasons and is also consistent with earlier studies of past European climates (e.g., Luterbacher et al., 2004; Brádzil et al., 2010; Büntgen 343 et al., 2011). However, between these two periods, multidecadal periods of cold anomalies 344 345 and warm anomalies, which are not similarly distributed in all seasons, can be identified (Fig. 3a). For instance, cold winter anomalies are observed up to the end of the 18th century: this is 346 followed by an increase in winter warm anomalies. In summer, multidecadal cold (19th 347 century) and warm periods (18th and 20th centuries) alternate (Fig. 3a). This therefore 348 349 highlights seasonal aspects of temperature fluctuations over the past centuries, which are also 350 consistent with earlier studies of past European climate. Here, the transition between the LIA and the Modern period has therefore been placed in the 19th century, *i.e.*, when winter 351 352 temperature begins to warm up, in accordance with other studies of European and Northern Hemisphere climate (e.g., Bradley & Jones, 1993; Jones et al., 1998; Moberg et al., 2005; 353 Ljungqvist, 2010; Büntgen et al., 2011). 354

355

We also identify significant temporal changes in precipitation over northern France. Measurements from the Paris Observatory show that the seasonal precipitation distribution

changes from one with a pronounced summer peak at the beginning of the record in 1688 358 until the 1750s, to an almost uniform seasonal distribution in the 20th century (Slonosky, 359 2002) (Fig. 3b). The seasonal reconstruction by Pauling et al. (2006) overestimates winter and 360 autumn precipitation over northern France (Figs. 3c, 4a, d). This is particularly true during the 361 first half of the 18th century, when precipitation is estimated using the Paris observations and 362 363 several English early-instrumental series as predictors (Pauling *et al.*, 2006). However, as in the Paris observations, winter and spring are clearly wetter during the 20th century than during 364 365 the previous two centuries (Fig. 3c).

366

The low winter values during part of the observed record (1688–1754) could, however, be due 367 to imperfect measurement of snow, although the notes published in the *Mémoires* indicate that 368 the snow was melted before being measured, which would give a reasonably accurate 369 370 measurement of the liquid water content. Evaporation or sublimation loss from melted or 371 solid snow, or reduced catch of snow due to changes in the instrument location may have contributed to the low winter values. However, to balance this, the exposed East Tower 372 should have acted as a wind shield and snow fence, increasing the catch. It was moreover 373 374 noted at the time by de la Hire (1712) that this seasonal distribution (Fig. 3), with the majority 375 of the precipitation falling in the summer months, was considered normal.

376

4.2 Changes in seasonal averages, and multiscale trend analysis

Changes in seasonal averages and trends in precipitation were investigated for both observed and reconstructed data, and results shown in Figure 4. Average winter and spring precipitation shows several positive shifts in the mean since the mid-18th century (Fig. 4a-b). Trend analysis, using the *Sen's* slope estimator and modified Mann-Kendall test of observed and reconstructed data, reveals that winter and spring precipitation show significant positive

secular trends at p = 0.05 since the mid-19th century, and even earlier for spring reconstruction 383 (Fig. 4e-f, i-j). Since the mid-19th century, these secular precipitation trends seem strongly 384 influenced by semi-secular periods of alternating increasing and decreasing trend, especially 385 in spring (Fig. 4e-f, i-j). Since the mid-18th century, average summer and autumn precipitation 386 trends are not significant, except during very wet years (Fig. 4c, d, g-h, k-l). Semi-secular 387 388 periods of alternating increasing and decreasing trends are also identified in summer precipitation before the 20th century or, for the reconstruction, before the mid-19th century 389 (Fig. 4g-k). This is less clear in autumn, and it appears only in reconstruction (Fig. 4h, 1). 390 391 These results are consistent with early instrumental rain-gauge records in England and 392 northern France (Dieppois, 2013), and also with historical precipitation reconstructions for Switzerland of Pfister (1994), and Alpine precipitation reconstructions of Auer et al. (2005) 393 and Casty et al. (2005). 394

395

We then examined whether such changes in the winter, spring and summer precipitation are 396 associated with modifications in temperature and PE (Fig. 5). Only winter temperature, and its 397 associated PE, shows a significant change in the mean in the mid-19th century (Fig. 5a), which 398 is related to significant warming trends since the beginning of the 19th century (Fig. 5d-i). 399 This trend in precipitation and temperature displays a significant positive correlation at semi-400 401 to multi-secular time-scales (Fig. 6a). According to Trenberth and Shea (2005), such a 402 relationship dominates when the water-holding capacity of the atmosphere limits precipitation 403 amounts in cold conditions, and warm air advection in cyclonic storms is accompanied by precipitation. For time-scales of several decades, however, this could also be consistent with 404 atmospheric circulation changes over the North Atlantic, such as produced in winter by the 405 NAO over northern Europe (Hurrell, 1995; Cassou et al., 2004; Hurrell and Deser, 2009; 406 Fleig et al., 2015). 407

408

For spring and summer, this trend and change in average temperature occurs much later, *i.e.*, 409 in the second half of the 20th century (Fig. 5b-c, e-f, j-k). Nevertheless, winter temperature 410 and precipitation are significantly correlated with spring precipitation at time-scales greater 411 412 than 50 years (Fig. 6b, d). Meanwhile, summer precipitation is negatively correlated to winter temperature from the beginning of the 19th century (Fig. 6c). The rising trend in winter 413 414 temperature and precipitation could therefore slightly influence the trend in spring 415 precipitation, potentially via a persistent response to a long-term change in the winter atmospheric circulation, at least since the 19th century. Such a relationship between winter and 416 417 spring hydroclimatic conditions might also involve the influence of soil moisture. According 418 to Schär et al. (1999) and Rowell and Jones (2006), wetter soils associated with wetter winters could lead to higher moist static energy per unit of planetary boundary layer air, and therefore 419 be favourable to convective precipitation in the following spring. Furthermore, secular trends 420 in temperature and PE since the mid-19th century are modulated by semi-secular periods of 421 422 alternating positive and negative trends in all seasons (Fig. 5 d-k). In winter and spring, these multidecadal periods in precipitation, temperature and PE are thus in-phase, *i.e.*, wet 423 conditions associated with warm anomalies and more potential evapotranspiration, which 424 425 could be consistent with NAO-like anomalies.

426

In summary, winter and spring trends in precipitation are detected starting from the end of the 19th century. Interestingly, the observations here contrast with changes in European precipitation seasonality revealed by a four-century-long-tree-ring isotopic record from Brittany in western France (Masson-Delmotte *et al.*, 2004), which suggested drier winters and wetter summers since the beginning of the 19th century. Additionally, this hydroclimatic trend could be modulated by semi-secular or multidecadal fluctuations. 433

434 **4.3 Multidecadal variability of seasonal averages**

Dieppois et al. (2013) showed that observed northern France precipitation demonstrated 435 436 significant variability on time-scales of 50–80 years, 16–23 years, 9–16 years and 2–8 years. In the present paper, we show that such time-scales of precipitation variability are also 437 438 characterized by different periods of increasing precipitation variance, which can reveal 439 seasonal differences in both observed and reconstructed data, especially at the multidecadal 440 (50–80 year) scale (Fig. 7a-f). These periods of increasing variance are statistically significant at p = 0.05. The 50–80 year variability of observed and reconstructed precipitation increases 441 slightly since the end of the 19th century in winter (Fig 7a-b) and, more significantly, in spring 442 (Fig. 7c-d). In summer, the variance of observed and reconstructed 50-80 year fluctuations 443 decreases throughout the 19th century (Fig. 7e, f). These seasonal changes in the multidecadal 444 variability of precipitation in northern France are also identified in early-instrumental rain-445 gauge records in England and northern France (Dieppois, 2013), and also with historical 446 precipitation reconstructions for Switzerland by Pfister (1994). 447

448

449 The wavelet spectra of average seasonal temperatures are displayed in Figure 8a-c (the same 450 time-scale patterns are observed in PE, not shown here). The 50–80 year fluctuations emerge 451 in the temperature series (Fig. 8a-c). Significant periods of high variance in 50-80 year 452 temperature variability are detected before 1750 in spring and summer, and in all seasons since the end of the 19th century (Fig. 8a-c). In winter and spring, increasing multidecadal 453 fluctuations in precipitation since the end of the 19th century could thus be associated with 454 multidecadal fluctuations in temperature (Figs. 7, 8). As mentioned in Section 4.2, an in-phase 455 456 relationship is identified, which suggests warm (cold) air advection in cyclonic storms, such as observed during positive (negative) NAO (Hurrell, 1995). Seasonal modifications of 50-80 457

year NAO variability have therefore been investigated using PC-based seasonal NAO indices 458 (Fig. 8d-f). Depending on the season, different periods of increasing variance are identified in 459 the multidecadal NAO variability (Fig. 8d-f). For instance, we identify increased multidecadal 460 variability of the winter and spring NAO from the end of the 19th century. Multidecadal 461 variability of the summer NAO significantly decreases throughout the 19th century in tandem 462 463 with multidecadal precipitation variability (Figs. 5, 7c, f). These multidecadal changes in the NAO indices at the end of the 19th century are also evident in the reconstructed fields of 464 Küttel et al. (2010) and in the ADVICE and HadSLP2r observations. 465

466

From the end of the 19th century, increased multidecadal variability of precipitation and 467 temperature is associated with similar changes in the summer and spring NAO. Multidecadal 468 fluctuations induced by atmospheric circulation changes over the North Atlantic could 469 therefore have substantially modulated the trends in the climatic variables over the 20th 470 century. Furthermore, as shown in Figure 6, winter temperature and precipitation are 471 positively correlated with spring precipitation at timescales greater than 50 years. This may 472 reflect a delayed response to the long-term change in the winter atmospheric circulation, 473 and/or soil moisture feedbacks (cf. Schär et al., 1999; Rowell and Jones, 2006). This might 474 therefore explain why multidecadal precipitation variability is more pronounced in spring. 475

476

477

478 **4.3. Impact of climate variations on Seine River flow**

As described in Section 3.2, the potential effects of secular and multidecadal changes in precipitation, temperature and PE on Seine river flow have been evaluated by hydrological modelling using GR2M, with the results displayed in Figure 9.

482

Simulated mean winter and spring river flows of the Seine at Paris show a positive shift at the 483 beginning of the 20th century (Fig. 9d-e). Trend analysis reveals a very good match with the 484 observations over the common period (*i.e.*, 1885–2009; Corr. \geq 0.868, significant at p = 0.05), 485 and shows significant and positive secular trends since the second half of the 19th century 486 (Fig. 9d-e). This is the first time that such positive trends have been identified in Seine river 487 488 flow: hitherto, these had been obscured by significant variability at the multidecadal timescale (Fig. 9d-e, g-h). The multidecadal timescale of variability does indeed reveal a period of 489 increasing variance in winter and spring since the end of the 19th century (Fig. 9g-h). This is 490 particularly pronounced in spring using both simulated and observed Seine River flow. Spring 491 492 is therefore an appropriate season to examine the effects of multidecadal fluctuations on the 493 trend. For instance, a negative shift in the mean occurs during the 1940s and 1950s when multidecadal modulations of the trend changes sign (*i.e.*, positive to negative trends; Fig. 9b, 494 495 e). In summer, simulated and observed river flows do not show any significant trends over the last centuries (Fig. 9c, f). Multidecadal fluctuations, which decrease throughout the 19th 496 century, are also evident (Fig. 9f, i). 497

498

Significantly, this modelling show that changes affecting trends and multidecadal variability in precipitation and temperature over the last centuries are also identifiable in Seine river flow. This is also consistent with the findings of Boé and Habets (2014) who first highlighted that multidecadal variability in river flows over France were more pronounced in spring.

503

504 5. Changes in multidecadal North Atlantic SLP variability

505 5.1. Multidecadal North Atlantic SLP anomalies and potential NAO linkages

Figure 10 shows the multidecadal anomalies (>30-years) of reconstructed North Atlantic SLPs compared to NAO patterns from the unfiltered HadSLP2r dataset, which were all captured by the first-EOFs of each season for the Modern period.

509

510 Observed NAO patterns consist of north-south dipole anomalies, with one centre of action 511 located over Greenland and the other centre of opposite sign spanning the subtropical latitudes 512 of the North Atlantic (Fig. 10a-c). These are more notable in winter (~47.1% of the total 513 variance). The subtropical centre is located between 35°N and 40°N in winter, but is displaced significantly north-eastward in spring and, especially, in summer (Hurrell and Deser, 2009; 514 515 Folland et al., 2009; Fig. 10a-c). In winter and spring, multidecadal anomalies of North 516 Atlantic SLPs also display north-south dipole anomalies (Fig. 10d-e, g-h, j-k, m-n). By 517 comparing the observed NAO patterns with multidecadal North Atlantic SLP anomalies, a 518 slight southward shift of the winter NAO center of action is observed over the eastern region 519 between Greenwich, London and 40°E (Fig. 10d, g, j, m). Such modifications are more pronounced in spring: a south-eastward (south-westward) shift of the Icelandic Low (Azores 520 521 High) occurs in all reconstructions and observations (Fig. 10e, h, k, n).

522

These modifications in the North Atlantic meridional SLP gradient could therefore alter the 523 524 zonal atmospheric circulation associated with the winter and spring NAO pattern. These 525 multidecadal anomalies are indeed spatially similar to those of the observed NAO, especially 526 in winter (sCorr. ≥ 0.90 , significant at p = 0.05), and, albeit with a reduce consistency 527 between datasets, in spring $(0.18 \ge sCorr. \ge 0.72)$, significant at p = 0.05). In addition, with the exception of ADVICE observations, these anomalies also show significant temporal 528 529 correlations at p = 0.05 with multidecadal fluctuations extracted from NAO indices (0.44 \geq *tCorr*. \geq 0.74). 530

During summer, with the exception of reconstructions by Luterbacher *et al.* (2002), multidecadal anomalies display a SLP pattern centred on northeastern Europe with SLP anomalies of opposite signs over the Atlantic and North Africa (Fig. 10f, i, 1, o). These multidecadal summer anomalies are thus unlikely to be related to the NAO ($0.01 \le sCorr. \le$ 0.31; $-0.05 \le tCorr. \le 0.34$). According to the CWT of NAO index (Fig. 8f), this describes a weakening in the summer NAO multidecadal variability during the Modern period (Fig. 10f, i, 1, o).

539

The multidecadal anomalies of North Atlantic SLPs reveal changes, which are likely to alter the NAO patterns, especially in winter and spring. Such multidecadal modulations of the NAO might be expected to be associated with changes in North Atlantic storm tracks. This will also influence northwestern European hydroclimatic conditions (cf. Section 4).

544

545 **5.3. Multidecadal correlation patterns**

By computing the correlation between 30-year running mean of precipitation and the multidecadal North Atlantic anomalies and comparing them with correlation patterns between unfiltered datasets (Fig. 11), we can now examine how North Atlantic multidecadal variability has affected seasonal relationships between the NAO and northwestern European precipitation.

551

The NAO patterns are associated with changes in the intensity and location of the North Atlantic jet stream and storm tracks; such large-scale modulations of the normal patterns of zonal and meridional moisture transport can modify precipitation patterns (Hurrell, 1995; Hurrell and Deser, 2009; Folland *et al.*, 2009). This is likely, then, to explain the statistically

⁵³¹

556 significant connection between the NAO and European precipitation during each season. The 557 correlation patterns are characterized by a strong north-south dipole in precipitation, which is associated with wet and dry conditions over northwest Europe and the Mediterranean region, 558 respectively, during the positive phase of the winter NAO (Fig. 11a). In spring and summer, 559 this north-south dipole in precipitation is shifted northeastwards in response to seasonal 560 561 displacements of NAO patterns (Fig. 11b-c). In all seasons, the positive phase of the NAO is usually negatively correlated with precipitation in northern France (Fig. 11a-c). This 562 563 connection, however, is more pronounced in summer, when anticyclonic conditions, which are displaced northeastward, suppress precipitation in northern France. 564

565

In winter and spring, there are statistically significant connections between the multidecadal 566 North Atlantic SLP anomalies, which are significantly linked to the NAO (Section 5.2), and 567 European precipitation in winter and spring (p-Field < 20% in reconstructions and 568 observations; Fig. 11d-e, g-h, j-k, m-n). However, during summer, with the exception of 569 ADVICE observations, the map is indistinguishable from random noise (*p*-Field > 20% in 570 reconstructions and observations; Fig. 11f, i, l, o). This multidecadal variability modulating 571 572 the winter NAO is associated with dipole anomalies in precipitation (Fig. 11d-e, g-h, j-k, mn). Positive anomalies of the multidecadal North Atlantic SLP variability lead to wet 573 574 conditions over Scandinavia, the Baltic countries, the British Isles and the Netherlands, while 575 they are associated with dry conditions over the Mediterranean regions. This is very similar to 576 precipitation anomalies induced by the observed NAO, but we note positive correlations (only 577 significant in the HadSLP observations) in northern France (Fig. 11d-e, g-h, j-k, m-n).

578

579 This suggests a southward shift in the winter NAO pattern, which might alter the westerly 580 storm tracks at the multidecadal time-scale. Such a shift in precipitation response to changes in the multidecadal North Atlantic SLP variability is more pronounced in spring (Fig. 11e, h, k, n), when a south-eastward (south-westward) shift of the Icelandic Low (Azores High) occurs in all reconstructions and observations. A substantial shift of the NAO pattern, which might alter the westerly tracks, is also suggested at the multidecadal time-scales for spring. In spring, the multidecadal correlation patterns are indeed regionally reversed compared to that of the spring NAO. Significant positive correlations are identified over France and centraleastern Europe (Fig. 11e, h, k, n).

588

589 6. Discussion and conclusion

590 This study has defined, through detailed statistical analysis, the magnitude and direction of 591 seasonal trends and shifts, over the last 500 years, in key hydroclimatic variables of northern France. It has also identified, especially through X and Y approaches, important potential 592 593 multidecadal modulations that can themselves be driven by large-scale atmospheric variability. The evidence presented here shows that substantial changes have occurred over 594 the past few centuries in the hydroclimate of northern France. In particular, several positive 595 shifts and an overall increasing trend are particularly noticeable in winter and spring 596 precipitation since the mid-19th century. This is linked quantitatively to rising winter 597 temperatures during the same period. 598

599

In addition, hydrological modelling has shown that hydroclimatic trends have helped to drive significant increases in mean Seine river flows at Paris, but only in winter and spring. Further work will address additional drivers of river flow variations over the past five centuries related to geomorphological changes in the watershed, and/or modifications in the catchment properties associated with land use changes, such as forest clearance, shifts to arable farming, increasing viticulture (especially in the Champagne region of eastern part of the Seine basin),

as well as social landholding fragmentation (dismemberment) processes which were very
common in France, but which are now being addressed through "rememberment" initiatives.
Such land use changes elsewhere have typically led to reductions in infiltration rates,
increased runoff, enhanced sediment losses from catchment slopes to river channels, and
declines in groundwater recruitment and storage.

611

Nevertheless, we have found that such secular trends in the seasonal hydroclimatic conditions over northern France are actually associated with key multidecadal fluctuations. Indeed, increasing variance of hydroclimatic multidecadal variability since the end of the 19th century has emerges strongly, but mainly/only in winter and spring seasons.

616

Such multidecadal hydrological fluctuations are driven mainly by large-scale circulation 617 618 changes. Similar connections have been previously revealed for eastern North American and 619 UK hydrosystems (Phillips et al., 2003; Dixon et al., 2006; Kingston et al., 2006a, b, 2009, 2011; Lavers et al., 2010; Holman et al., 2011). However, these have been for much shorter 620 621 periods than the 500 years here. Also the more complex and exhaustive analyses here have 622 substantially extended understanding of these hydroclimatic fluctuations and their circulation drivers into a new key region of North Atlantic, Europe and northern France (cf. Massei et al., 623 624 2007).

625

Our results also demonstrate how North Atlantic multidecadal variability, which arises principally from internally generated natural processes, are likely to influence recent trends (or lack thereof) in hydroclimatic variables, and which are very likely to contribute to future climate variability (Keenlyside *et al.*, 2008; Watanabe *et al.*, 2014). Multidecadal NAO variability has increased since the 19th century in winter and spring. 631

Furthermore, multidecadal anomalies of North Atlantic SLPs since the 19th century reveal changes in the meridional dipole, which are statistically linked to modulations of the NAO. In particular, in spring, a south-eastward (south-westward) shift of the Icelandic Low (Azores High) drives significant multidecadal hydroclimatic changes: this leads to modulations of moisture advection in response to shifts of North Atlantic storm tracks over European regions, and wetter springs in northern France during positive NAO phases. Similar changes are observed in winter, although these are less pronounced.

639

Recently, based on the results from CCSM4 model, Sun et al. (2015) proposed two possible 640 641 mechanisms responsible for the multidecadal variability of the NAO: the forcing of the Atlantic Multidecadal Oscillation (AMO; Knight et al., 2005) by the NAO, and delayed 642 negative feedback exerted by the AMO. At the same time, as proposed by Franckombe et al. 643 (2010), the North-Atlantic multidecadal variability is also related to exchange processes 644 between the Atlantic and Arctic Oceans. This idea is consistent with the work of Miles et al. 645 (2014) who identify concomitant reinforcement of multidecadal variability in the AMO and 646 several Arctic sea-ice indices at the end of the 19th century. Also, it is possible that anomalous 647 sea-ice and freshwater pulses through Fram Strait could have helped to drive, as presented 648 649 here, an eastward shift in the Iceland Low over recent decades (Hilmer and Jung, 2000). 650 Intriguingly, this could thus form a hitherto unrecognized but fundamental and interactive 651 component of multidecadal North Atlantic variability in the post-LIA period, and consequently affect multidecadal variability of hydrosystems in France and over NW Europe 652 653 more generally. This requires testing in future work.

654

The possible influence of an early warming winter temperature over the Arctic regions on 655 Arctic sea-ice export, which has been detected here in central England and Paris (not shown), 656 should be explored very carefully (see, e.g., Slonosky et al. (1997) and Mysak et al. (1996) 657 for multidecadal sea-ice and atmospheric circulation anomalies in the 20th century). This 658 might have played a role in initiating increased multidecadal climate variability over the 659 North Atlantic. External forcing could also be involved in the relationship between North 660 Atlantic multidecadal variability and European hydroclimatic variability. Recently, Knudsen 661 662 et al. (2014) suggested that external forcing, *i.e.*, solar variability or volcanic forcing, played a dominant role in pacing the AMO after termination of the LIA. Also, anthropogenic and 663 664 natural aerosols have likely played some role in forcing the Atlantic multidecadal variability (Chang et al., 2011; Booth et al., 2012; Villarini and Vecchi, 2013). 665

666

Our study also demonstrates that warming trends began earlier than hitherto recognised, 667 which are associated with changes in PE, precipitation and, potentially, in river flow in a 668 major European hydroclimatic system (northern France and the Seine river basin). Until now, 669 670 these trends were masked by aperiodic multidecadal oscillations that originated from internal 671 variability in the North Atlantic climate system. Although this could involve non-linear and chaotic interactions, which are not yet well understood, the impact of such multidecadal 672 oscillations on hydroclimate in northern France seems as important as the secular trends since 673 the mid-19th century. Understanding future regional and continental hydroclimatic trends, 674 675 especially precipitation and river flow over the next 50 years is a challenge. However, it is likely that illumination of this key question could be driven by improved knowledge of the 676 links between North Atlantic multidecadal climatic variability and hydrological impacts over 677 678 the European continent.

679

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- 682

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944 945 946 **List of Figures** 947 Figure 1. Overview of the data sets used in this study against time-scale. Observation and 948 949 reconstruction data are indicated in blue and orange, respectively. Time series and gridded data are displayed in lines and boxes, respectively. Time series derived from calculations are 950 951 in red. 952 Figure 2. Spatial coherence of each seasonal precipitation time series. (a-d) Seasonal 953 954 pointwise correlations precipitation anomalies of observed Paris rain-gauges and the CRU TS 3.10.1 field over the NW Europe (top: DJF; top center: MAM, bottom center: JJA, bottom: 955 SON). (e-h) As in (a-d) using a reconstructed northern France precipitation index (47.5 -956 957 50.5°N; 0 - 4°E) and the high-resolution precipitation reconstruction of Pauling *et al.* (2006). 958 Red contours indicate correlations significant at the 95% confidence level (p < 5%). 959 Figure 3. Temporal evolution of seasonal distributions of temperature and precipitation in 960 northern France. (a) Observed secular distributions of monthly Central England Temperature 961 anomalies (°C) with regard to the 20th century mean (red: warm anomalies; blue: cold 962 anomalies). (b) Observed secular distributions of monthly Paris precipitation (mm.month⁻¹) 963 since 1700. (c) As in (b) for a seasonal reconstructed northern France precipitation index 964 (mm.season⁻¹) since 1500. Grey contours delineate the area above which precipitation is 965 greater than the multi-decadal mean (~60 year smoothing window). 966

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Figure 4. Modifications in mean and trend of seasonal precipitation amounts in northern 969 970 France over the last five centuries. (a-d) Multiples changes of the means in the observed (grey) and reconstructed (light red) seasonal precipitation time series (mm.season⁻¹; winter: 971 DJF; spring: MAM; summer: JJA; autumn: SON). Bold lines display optimal breaks between 972 adjacent segments with means that are significantly different from the Scheffe contrast test at 973 974 the 95% confidence interval as determined by a segmentation procedure. (e-h) Two-975 dimensional diagrams of every possible trend in the observed seasonal precipitation time 976 series as determined by a modified Mann-kendall test. Trend strength (positive: blue; negative: red) is estimate using the Sen's slope while statistical significance at p = 0.05977 978 (contour) is assumed by the two-sided *p-values*. (i-l) As in (e-h) for a seasonal reconstructed northern France precipitation index (mm.season⁻¹) since 1500. 979

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981 Figure 5. Modifications in mean and trend of seasonal central England temperature and its 982 associated evapotranspiration (PE) over the last centuries. (a-c) Multiples changes of the means in observed temperature (°C; grey) and PE (mm; light red) winter: DJF; spring: MAM; 983 summer: JJA). Bold lines display optimal breaks between adjacent segments with means that 984 are significantly different from the Scheffe contrast test at the 95% confidence interval as 985 determined by a segmentation procedure. (d-f) Two-dimensional diagrams of every possible 986 987 trend in observed temperature as determined by a modified Mann-kendall test. Trend strength 988 (positive: red; negative: blue) is estimate using the Sen's slope while statistical significance at p = 0.05 (contour) is assumed by the two-sided *p*-values. (g-k) As in (e-h) for PE (positive: 989 blue; negative: red) since 1500. 990

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Figure 6. Time-scale evolution of relationships between winter temperature/precipitation and 994 995 seasonal precipitation average. (a-c) Two-dimensional diagrams of every possible correlation between winter temperature and winter precipitation (at the top), spring precipitation (at the 996 997 middle) and summer precipitation (at the bottom). (d-e) as for (b-c) but for the correlation between winter precipitation and spring and summer precipitation. Contours indicate 998 999 correlations statistically significant at p = 0.05 of Pearson's product moment correlation coefficient assuming independent normal distributions. Note that the degrees of freedom 1000 1001 (DOF) in the significance calculations are adjusted using the estimated decorrelation scales.

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Figure 7. Time-scale patterns of variability in the seasonal precipitation amounts of northern 1004 France. (a-f) Wavelet power spectrum of observed (left column) and reconstructed (right 1005 column) seasonal precipitation time series (winter: DJF; spring: MAM; summer: JJA). The 1006 very thick bold lines (the so-called cone of influence) delineates the area under which power 1007 can be underestimated as a consequence of edge effects, wraparound effects and zero 1008 padding; thin contour lines show the 95% confidence limits based on 1000 Monte-Carlo 1009 1010 simulations of the red noise background spectrum. Note that these time-scale patterns were 1011 also obtained without padded with zeroes near the edge of the time series.

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1018	Figure 8. Time-scale patterns of variability in seasonal temperature (and PE) and in the NAO
1019	indices. (a-c) Wavelet power spectra of seasonal central England temperature (CET; °C),
1020	which are similar to those of PE (winter: DJF; spring: MAM; summer: JJA). (d-f) as in (a-c)
1021	for the PC-based seasonal NAO indices. The very thick bold lines (the so-called cone of
1022	influence) delineates the area under which power can be underestimated as a consequence of
1023	edge effects, wraparound effects and zero padding; thin contour lines show the 95%
1024	confidence limits based on 1000 Monte-Carlo simulations of the red noise background
1025	spectrum. Note that these time-scale patterns were also obtained without padded with zeroes
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Figure 9. Evaluation of modifications in mean, trend and decadal variability of Paris river 1042 flow by GR2M hydrological modelling. (a-c) Multiples changes of the seasonal means 1043 (winter: DJF; spring: MAM; summer: JJA) in observed (grey lines) and simulated Paris river 1044 flow (red lines; m³.s⁻¹). Bold lines display optimal breaks between adjacent segments with 1045 means that are significantly different from the Scheffe contrast test at the 95% confidence 1046 1047 interval as determined by a segmentation procedure. The light blue boxes indicate goodness-1048 of-fit measures between observed and simulated values with regard to the common period 1049 (1885–2009) before (black) and after (grey) the mean corrections. (d-f) Two-dimensional diagrams of every possible trend in seasonal observed (small box) and simulated river flow as 1050 1051 determined by a modified Mann-kendall test. Trend strength (positive: blue; negative: red) is 1052 estimate using the Sen's slope while statistical significance at p = 0.05 (contour) is assumed by the two-sided *p-values*. Correlations between observed and simulated trend patterns, and 1053 their statistical significances at p = 0.05, over the common period are indicated on the upper 1054 right corners of the observed diagrams. (g-i) Wavelet power spectra of seasonal observed 1055 (small box) and simulated river flow. The very thick bold lines (the so-called cone of 1056 influence) delineates the area under which power can be underestimated as a consequence of 1057 edge effects, wraparound effects and zero padding; thin contour lines show the 95% 1058 confidence limits based on 1000 Monte-Carlo simulations of the red noise background 1059 spectrum. 1060 1061

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Figure 10. Seasonal NAO patterns and multidecadal anomalies in North Atlantic SLP during the Modern period. (a-c) Seasonal 1^{st} EOF of unfiltered North Atlantic SLP ($30^{\circ} - 70^{\circ}N$; 30°W – 40°E), *i.e.*, NAO patterns, from the HadSLP2r dataset during the modern period. (d-f) Reconstruction using EOF of multidecadal anomalies (i.e., 30-year running mean) in the winter to summer observed North Atlantic SLP during the modern period. (g-l) as for (d-f) but using SLP reconstructions from Luterbacher et al. (2002) and Küttel et al. (2010). (m-o) idem based on ADVICE observation fields. Fraction of the variance expressed by each 1st EOFs are displayed on the upper right corners. Spatial and temporal correlation (sCorr. and tCorr.) between multidecadal anomalies and the NAO (i.e., observed NAO patterns and multidecadal variability extracted from the NAO indices for each dataset) are indicated on the lower left corners. Asterisks indicate significant correlations at p = 0.05 with regard to Dutilleul's and Pearson's *t*-tests accounting for spatial and temporal autocorrelation.

Figure 11. Seasonal correlation patterns between European precipitation and the NAO, as well 1092 as the multidecadal North Atlantic SLP anomalies during the Modern period. (a-c) Pointwise 1093 correlations between seasonal PC-based NAO indices from HadSLP2r dataset and European 1094 1095 precipitation during the Modern period. (d-f) Seasonal pointwise correlations between the 30year running mean of European precipitation and observed multidecadal North Atlantic SLP 1096 1097 anomalies during the Modern period. (g-l) as for (d-f) but using SLP reconstructions from 1098 Luterbacher et al. (2002) and Küttel et al. (2010). (m-o) idem based on ADVICE observation 1099 fields. Note that the degrees of freedom (DOF) in the significance calculations are adjusted using the estimated decorrelation scales, so the signal may not be more significant after the 1100 low-pass filtering. The p-values at the 95% (i.e., p=0.05) confidence level are computed 1101 against a 1000 sample Monte-Carlo simulations, where the precipitation fields are replaced 1102 β Spu. with red noise. Field significance and the spatial DOF are displayed on the upper right 1103 1104 corners.



Figure 1. Overview of the data sets used in this study against time-scale. Observation and reconstruction data are indicated in blue and orange, respectively. Time series and gridded data are displayed in lines and boxes, respectively. Time series derived from calculations are in red. 219x263mm (300 x 300 DPI)



Figure 2. Spatial coherence of each seasonal precipitation time series. (a-d) Seasonal pointwise correlations precipitation anomalies of observed Paris rain-gauges and the CRU TS 3.10.1 field over the NW Europe (top: DJF; top center: MAM, bottom center: JJA, bottom: SON). (e-h) As in (a-d) using a reconstructed northern France precipitation index (47.5 – 50.5°N; 0 – 4°E) and the high-resolution precipitation reconstruction of Pauling et al. (2006). Red contours indicate correlations significant at the 95% confidence level (p<5%). 444x598mm (300 x 300 DPI)

Figure 3. Temporal evolution of seasonal distributions of temperature and precipitation in northern France.
 (a) Observed secular distributions of monthly Central England Temperature anomalies (°C) with regard to the 20th century mean (red: warm anomalies; blue: cold anomalies).
 (b) Observed secular distributions of monthly Paris precipitation (mm.month-1) since 1700.
 (c) As in (b) for a seasonal reconstructed northern France precipitation index (mm.season-1) since 1500. Grey contours delineate the area above which precipitation is greater than the multi-decadal mean (~60 year smoothing window).
 291x298mm (300 x 300 DPI)

Figure 4. Modifications in mean and trend of seasonal precipitation amounts in northern France over the last five centuries. (a-d) Multiples changes of the means in the observed (grey) and reconstructed (light red) seasonal precipitation time series (mm.season-1; winter: DJF; spring: MAM; summer: JJA; autumn: SON). Bold lines display optimal breaks between adjacent segments with means that are significantly different from the Scheffe contrast test at the 95% confidence interval as determined by a segmentation procedure. (e-h) Two-dimensional diagrams of every possible trend in the observed seasonal precipitation time series as determined by a modified Mann-kendall test. Trend strength (positive: blue; negative: red) is estimate using the Sen's slope while statistical significance at p = 0.05 (contour) is assumed by the two-sided p-values. (i-l) As in (e-h) for a seasonal reconstructed northern France precipitation index (mm.season-1) since 1500.

187x104mm (300 x 300 DPI)

Figure 5. Modifications in mean and trend of seasonal central England temperature and its associated evapotranspiration (PE) over the last centuries. (a-c) Multiples changes of the means in observed temperature (°C; grey) and PE (mm; light red) winter: DJF; spring: MAM; summer: JJA). Bold lines display optimal breaks between adjacent segments with means that are significantly different from the Scheffe contrast test at the 95% confidence interval as determined by a segmentation procedure. (d-f) Two-dimensional diagrams of every possible trend in observed temperature as determined by a modified Mann-kendall test. Trend strength (positive: red; negative: blue) is estimate using the Sen's slope while statistical significance at p = 0.05 (contour) is assumed by the two-sided p-values. (g-k) As in (e-h) for PE (positive: blue; negative: red) since 1500.

137x56mm (300 x 300 DPI)

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Figure 6. Time-scale evolution of relationships between winter temperature/precipitation and seasonal precipitation average. (a-c) Two-dimensional diagrams of every possible correlation between winter temperature and winter precipitation (at the top), spring precipitation (at the middle) and summer precipitation (at the bottom). (d-e) as for (b-c) but for the correlation between winter precipitation and spring and summer precipitation. Contours indicate correlations statistically significant at p = 0.05 of Pearson's product moment correlation coefficient assuming independent normal distributions. Note that the degrees of freedom (DOF) in the significance calculations are adjusted using the estimated decorrelation scales.

217x166mm (300 x 300 DPI)

Figure 7. Time-scale patterns of variability in the seasonal precipitation amounts of northern France. (a-f)
 Wavelet power spectrum of observed (left column) and reconstructed (right column) seasonal precipitation time series (winter: DJF; spring: MAM; summer: JJA). The very thick bold lines (the so-called cone of influence) delineates the area under which power can be underestimated as a consequence of edge effects, wraparound effects and zero padding; thin contour lines show the 95% confidence limits based on 1000
 Monte-Carlo simulations of the red noise background spectrum. Note that these time-scale patterns were also obtained without padded with zeroes near the edge of the time series. 269x255mm (300 x 300 DPI)

Figure 8. Time-scale patterns of variability in seasonal temperature (and PE) and in the NAO indices. (a-c) Wavelet power spectra of seasonal central England temperature (CET; °C), which are similar to those of PE (winter: DJF; spring: MAM; summer: JJA). (d-f) as in (a-c) for the PC-based seasonal NAO indices. The very thick bold lines (the so-called cone of influence) delineates the area under which power can be underestimated as a consequence of edge effects, wraparound effects and zero padding; thin contour lines show the 95% confidence limits based on 1000 Monte-Carlo simulations of the red noise background spectrum. Note that these time-scale patterns were also obtained without padded with zeroes near the edge of the time series.

268x253mm (300 x 300 DPI)

Figure 9. Evaluation of modifications in mean, trend and decadal variability of Paris river flow by GR2M hydrological modelling. (a-c) Multiples changes of the seasonal means (winter: DJF; spring: MAM; summer: JJA) in observed (grey lines) and simulated Paris river flow (red lines; m3.s-1). Bold lines display optimal breaks between adjacent segments with means that are significantly different from the Scheffe contrast test at the 95% confidence interval as determined by a segmentation procedure. The light blue boxes indicate goodness-of-fit measures between observed and simulated values with regard to the common period (1885-2009) before (black) and after (grey) the mean corrections. (d-f) Two-dimensional diagrams of every possible trend in seasonal observed (small box) and simulated river flow as determined by a modified Mannkendall test. Trend strength (positive: blue; negative: red) is estimate using the Sen's slope while statistical significance at p = 0.05 (contour) is assumed by the two-sided p-values. Correlations between observed and simulated trend patterns, and their statistical significances at p = 0.05, over the common period are indicated on the upper right corners of the observed diagrams. (g-i) Wavelet power spectra of seasonal observed (small box) and simulated river flow. The very thick bold lines (the so-called cone of influence) delineates the area under which power can be underestimated as a consequence of edge effects, wraparound effects and zero padding; thin contour lines show the 95% confidence limits based on 1000 Monte-Carlo simulations of the red noise background spectrum.

214x133mm (300 x 300 DPI)

Figure 10. Seasonal NAO patterns and multidecadal anomalies in North Atlantic SLP during the Modern period. (a-c) Seasonal 1st EOF of unfiltered North Atlantic SLP ($30^\circ - 70^\circ$ N; 30° W - 40° E), i.e., NAO patterns, from the HadSLP2r dataset during the modern period. (d-f) Reconstruction using EOF of multidecadal anomalies (i.e., 30-year running mean) in the winter to summer observed North Atlantic SLP during the modern period. (g-l) as for (d-f) but using SLP reconstructions from Luterbacher et al. (2002) and Küttel et al. (2010). (m-o) idem based on ADVICE observation fields. Fraction of the variance expressed by each 1st EOFs are displayed on the upper right corners. Spatial and temporal correlation (sCorr. and tCorr.) between multidecadal anomalies and the NAO (i.e., observed NAO patterns and multidecadal variability extracted from the NAO indices for each dataset) are indicated on the lower left corners. Asterisks indicate significant correlations at p = 0.05 with regard to Dutilleul's and Pearson's t-tests accounting for spatial and temporal autocorrelation. 391x403mm (300×300 DPI)

Figure 11. Seasonal correlation patterns between European precipitation and the NAO, as well as the multidecadal North Atlantic SLP anomalies during the Modern period. (a-c) Pointwise correlations between seasonal PC-based NAO indices from HadSLP2r dataset and European precipitation during the Modern period. (d-f) Seasonal pointwise correlations between the 30-year running mean of European precipitation and observed multidecadal North Atlantic SLP anomalies during the Modern period. (g-l) as for (d-f) but using SLP reconstructions from Luterbacher et al. (2002) and Küttel et al. (2010). (m-o) idem based on ADVICE observation fields. Note that the degrees of freedom (DOF) in the significance calculations are adjusted using the estimated decorrelation scales, so the signal may not be more significant after the low-pass filtering. The p-values at the 95% (i.e., p=0.05) confidence level are computed against a 1000 sample Monte-Carlo simulations, where the precipitation fields are replaced with red noise. Field significance and the spatial DOF are displayed on the upper right corners.

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