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The Impact of FDI on Nigeria's Exports: A Sectoral Analysis

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Abstract

Purpose – The purpose of this paper is to examine the FDI-exports relationship in Nigeria using disaggregated FDI and export data.

Design/methodology/approach – This paper applies the ARDL cointegration approach in examining the long-run relationship between FDI and exports.

Findings – Our results suggest that aggregate FDI has a positive and statistically significant long-run impact on total exports. Once exports are disaggregated into oil and non-oil exports, the positive, cointegrating relationship holds only for oil exports. When disaggregated by sector, primary sector and manufacturing sector FDI have a positive and significant long-run relationship with both total exports and oil exports but service sector FDI does not appear to have any significant influence on Nigerian exports.

Originality/value – This is the first paper that employs both sectoral FDI and disaggregated export data to examine the FDI-exports nexus in Nigeria.

Keyword FDI, Export, Nigeria

Paper type Research paper

1. Introduction

Although much empirical work has focused on understanding how foreign direct investment (FDI) and the underlying mediating factors affect the growth and productivity of a host country, the relationship between FDI and a host country's export performance has received considerably less attention. More so, many of the studies on the FDI-exports nexus have focused on developed and transition economies, with little attention being paid to developing economies, especially Sub-Saharan African (SSA) countries.

One of the earliest theoretical frameworks on the relation between FDI and international trade was advanced by Mundell (1957) in his seminal work based on the neoclassical Heckscher-Ohlin's model based on two countries, two commodities and two factors. In his analysis, Mundell relaxes the assumption of immobility of factors of production and argues that factor movement can substitute for trade. He hypothesises that trade barriers will encourage cross-border factor movement, and that trade, on the other hand, would boom in the presence of restrictions on factor movement. This implies that FDI and exports can substitute each other depending on the degree of openness of the host and home country. Later theories such as the internalisation theory (Buckley and Casson, 1976) further emphasise the substitutive relationship between FDI and trade. In deciding to enter a foreign market, firms are likely to undertake FDI rather than exports if the benefits that accrue to them by internalising their ownership advantages exceed the cost of market imperfections.

Although many of the theoretical treatments of FDI and trade have often included outward FDI and imports, the precise relationship between inward FDI and exports remains ambiguous and may depend on the type of FDI. According to Vuksic (2005), FDI that is predicated on taking advantage of the availability of natural resources or low-cost labour is more likely to directly promote exports. Vertically integrated FDI that is resource seeking

would be expected to increase the volume of the host country's exports, with the subsidiaries of the MNEs focused on exporting raw materials or intermediate products to their parent firm or other subsidiaries. The effect of horizontal FDI, on the other hand, on the export performance of a host country is ambiguous. Jensen (2002) observed that this type of FDI, especially when it is 'market-seeking', may not have any direct impacts on exports as it is targeted primarily at the host country market. However, Franco (2013) argues that market-seeking horizontal FDI does indeed have the potential to promote host country exports through export spillover to domestic firms. Her results show that while market-seeking FDI may not contribute directly to exports, by strengthening the links with domestic firms, it may indirectly boost the exporting capacity of domestic firms. In other words, market-seeking horizontal FDI also has a potentially positive effect on exports when it boosts local entrepreneurship through competition with domestic firms as well as through links with firms in their downstream and upstream of the MNE's supply chain.

The empirical literature on the effect of FDI on host country exports can be broadly divided into two strands: (i) studies that focus on the direct contribution of FDI to exports; and (ii) studies that examine the indirect (spillover) effect of FDI on domestic firms' exports. In the present study, due to a lack of firm-level data for Nigeria, we focus on the direct effect of FDI. The empirical evidence from this stream of literature is mixed. Some studies found evidence of a positive relationship (e.g., Leichenko *et al.*, 1997; Xuan and Xing, 2008), while others show that the impact of FDI on host country exports may not always be positive (e.g., Lall and Mohammed, 1983; Sharma, 2003). Furthermore, most of these studies make use of aggregate FDI and aggregate export data at either cross-national, national or regional level. As noted by Alfaro (2003), aggregate FDI does not give the full picture of the impact of FDI. Data aggregation may also be the reason for the mixed findings of previous work.

The present study adds to this literature by employing disaggregated sectoral FDI and oil/non-oil export data to examine the FDI-exports relationship in Nigeria. Nigeria makes for an interesting case to study the FDI-export nexus. For over a decade now, Nigeria has consistently remained one of the top destinations of FDI in Africa. Figure 1 shows the top 10 destinations of FDI in Africa (average from 1980 to 2015) and Nigeria is the second-ranked African country in terms of inward FDI. Moving away from a largely agrarian economy to an oil-dominated economy, Nigeria's FDI has also historically tilted towards the oil sector. As seen in Figure 2, for most of the 1990s, the primary sector (consisting largely of oil and allied sector) was the major recipient of FDI. However, this trend has been reversing since the early 2000s, as Nigeria has recorded massive FDI inflows into other sectors, particularly the manufacturing sector – also thanks to the extensive privatisation of public enterprises carried out by the new civilian government (UNCTAD, 2009).

[Insert Figure 1 and Figure 2 here]

Despite being one of the largest recipients of FDI in Africa, only very few empirical studies have examined the FDI-exports relationship in Nigeria. The notable exceptions are Mohammed and Ekundayo (2014), Aigheyisi (2016), and Enimola (2011). Unlike Enimola (2011), who investigated the FDI-export nexus at aggregate level, Mohammed and Ekundayo (2014) and Aigheyisi (2016) disaggregated export categories into oil and non-oil, taking into account the predominance of the oil sector in Nigeria. However, none of these studies attempted to examine the impact of sectoral FDI across different export categories.

We disaggregate FDI into primary sector, manufacturing sector FDI and service sector FDI, and investigate their impact across different export categories by distinguishing between oil and non-oil exports. By disaggregating FDI by sector, we are also able to infer how the 'FDI motivation' affects its relation with exports. Primary sector FDI, especially in

developing countries, is mostly resource-seeking and vertically integrated (Alfaro, 2003), and hence is more likely to promote resource exports (in this study, oil) than non-resource (non-oil) exports. Service sector FDI and manufacturing FDI, on the other hand, may either be market-seeking or efficiency-seeking, hence they may have a differentiated impact on resource and non-resource product exports.

2. A synthesis of related empirical literature

The empirical evidence is mixed. Studies by Leichenko *et al.* (1997), Sun (2001), Zhang (2002; 2005), Liu and Shu (2003), Wang *et al.* (2007) and Xuan and Xing (2008) find evidence of a positive relationship between FDI and exports at the national and regional level. However, there is also conflicting evidence for some countries. For example, Sharma (2003) investigated the contribution of FDI to India's exports over the period 1970-1998 and found that FDI did not have any significant impact. This result supports the earlier finding by Lall and Mohammed (1983) for India. To explain the absence or, at best, marginal contribution of FDI to exports in their findings, Lall and Mohammed (1983) and Sharma (2003) point out that the inward-oriented policy of the Indian government during the sample period, effectively discouraged exports. Likewise, in a study of 12 Central and Eastern European countries, Kutan and Vuksic (2007) found that the FDI-specific effect on exports was positive and statistically significant only for the eight countries grouped under the New European Union (NEU) members, and statistically insignificant for the other four Southeast European countries. They argue that the difference in FDI impact is likely to result from the fact that the level of FDI inflows and initial level of productivity of domestic firms in the NEU countries, are higher than in the Southeast European countries.

Another possible reason for the divergence in results of FDI-exports studies is that the type and level of aggregation of the data matters. Most existing studies make use of aggregate FDI and aggregate export data. However, it is likely that the use of aggregate FDI and export data masks sectoral differences in the impact of FDI on different export categories (Alfaro, 2003). Her results show that while total FDI may have a positive effect on the host country's growth, when disaggregated by sector, primary sector, manufacturing sector, and service sector FDI have different effects on growth. Similarly, while FDI inflows may promote exports at aggregate national or regional level, not all product exports may be affected in the same way. Moreover, although not related to the FDI-exports literature, Li *et al.* (2017) attribute the use of sectoral FDI to the underlying sector-specific characteristics and the fixed costs associated with the nature of such inward investments, which are generally higher for primary sector FDI. In turn, this would imply that oil exports which typically entail use of technological advantages embedded in 'resource seeking' inward FDI which is expected to incur high fixed costs, need to generate more profits through export revenues. Manufacturing and services FDI, on the other hand, incur lower fixed costs, are relatively more mobile, and hence less inclined to seek extra profits through exports.

In a study on the effect of FDI on the exports of 10 Economic Community of West African States (ECOWAS), Onyekwena *et al.* (2015) disaggregate exports into three categories: primary, intermediate, and final commodities. Their results show that FDI has a positive relationship with primary goods exports, a negative relationship with intermediate goods exports, and an insignificant impact on final goods exports. Using a similar approach, Gu *et al.* (2008) and Liu and Shu (2003) examine the impact of FDI on different categories of exports in China. Liu and Shu (2003) found that the effect of FDI on total exports is positive, and this result is robust when exports are disaggregated into high and low technology sectors' exports. Gu *et al.* (2008) obtained similar results after disaggregating industrial exports into

14 sectors. They find that the effect of FDI is positive and significant in 13 out of the 14 sectors considered. These findings are consistent with those of Leichenko and Erickson (1997), who found that the impact of FDI on exports is positive in all but two of the manufacturing industries considered. Taken together, what these studies suggest is that the impact of FDI (total or disaggregated) may not be the same across all export categories.

3. Theoretical and empirical framework

3.1 Hypotheses

We state two hypotheses to frame the empirical analysis.

Hypothesis 1: Aggregate inward FDI has a positive effect on the volume of exports in Nigeria.

As we have argued in the literature review, MNEs' affiliates are usually more productive and more likely to export as they possess specific ownership advantages, which may include having more knowledge about the workings of the international market. Moreover, MNEs are typically large firms, with superior technologies and financial means, hence better able than most domestic firms to afford the high fixed costs associated with exporting (Gorg and Greenaway, 2001). Hence, it is reasonable to expect that aggregate inward FDI may affect positively the volume of total exports in the host country.

While we expect that FDI has a positive effect on the total volume of exports, there is evidence in the literature that suggests that the impact of aggregate FDI on exports may vary across different export categories. For instance, Onyekwena *et al.* (2015) found that FDI does not have the same impact on primary products, intermediate and finished product exports in West Africa. This highlights the need to disaggregate exports into different categories in

order to assess the differential effect of aggregate FDI across the categories. Following related literature pertaining to the Nigerian experience, we disaggregate exports into oil and non-oil exports. The distinction is motivated by the fact that oil plays a significant role in Nigeria's economy. A large proportion of Nigeria's inward FDI goes to the oil sector and the sector's exports make up a very high percentage of total exports (Olayiwola and Okodua, 2013). Figure 3 shows the summary of exports in Nigeria and the dominance of oil exports in total exports. Consequently, this hypothesis aims to examine the effect of aggregate inward FDI on total exports, as well as oil and non-oil exports.

[Insert Figure 3 here]

Hypothesis 2: Sectoral FDI has a varied effect on export volume in Nigeria.

The aim of *Hypothesis 2* is to investigate the relationship between sectoral FDI - primary sector FDI, manufacturing sector FDI and service sector FDI - and Nigeria's exports. As noted in our literature review, increasing evidence suggests that the type and extent of the impact that FDI has on a host country will vary when total FDI is disaggregated by sectors (e.g., Alfaro, 2003; Li *et al.*, 2017). The use of aggregate FDI data may not reveal the whole picture, as it may obscure the fact that the impact of FDI is unlikely to be equal for all recipient sectors.

3.2 Model specification

In the previous section, we stated the two hypotheses to be subjected to empirical scrutiny.

To test the first hypothesis, we specify the following three models:

$$\ln EXP_t = \beta_0 + \beta_1 \ln FDI_t + \beta_2 \ln REER_t + \beta_3 \ln GDP_t + \mu_t \quad (1)$$

$$LnOEXP_t = \alpha_0 + \alpha_1 lnFDI_t + \alpha_2 lnREER_t + \alpha_3 lnGDP_t + \varepsilon_t \quad (2)$$

$$LnNEXP_t = \delta_0 + \delta_1 lnFDI_t + \delta_2 lnREER_t + \delta_3 lnGDP_t + \eta_t \quad (3)$$

In equations 1-3 above, the explanatory variables are the same while the dependent variable is different in each equation. *LnEXP* is the total amount of annual exports in log form, *LnOEXP* and *LnNEXP* stand for oil and non-oil exports in log form, respectively. *LnFDI*, *LnREER* and *LnGDP* represent aggregate FDI, the real effective exchange rate and gross domestic product, respectively, in log form.

Next, we propose three further models to test the second hypothesis. We disaggregate FDI into three broad sectors, primary sector FDI, manufacturing sector FDI, and service sector FDI, and estimate their long-run relationship with the different export categories.

$$LnTEXPt = n_0 + n_1 lnPFDI_t + n_2 lnMFDI_t + n_3 lnSFDI_t + n_4 lnREER_t + n_5 lnGDP_t + v_t \quad (4)$$

$$LnOEXP_t = h_0 + h_1 lnPFDI_t + h_2 lnMFDI_t + h_3 lnSFDI_t + h_4 lnREER_t + h_5 lnGDP_t + \varphi_t \quad (5)$$

$$LnNEXP_t = m_0 + m_1 lnPFDI_t + m_2 lnMFDI_t + m_3 lnSFDI_t + m_4 lnREER_t + m_5 lnGDP_t + \chi_t \quad (6)$$

The dependent and explanatory variables in equations 4-6 are identical to those in equations 1-3. The only difference is that, in equations 4-6, FDI is disaggregated into primary sector FDI (*PFDI*), manufacturing sector FDI (*MFDI*), and service sector FDI (*SFDI*).

3.3 Data and method

Annual time series data were obtained from different sources. Data on FDI, REER and GDP were obtained from the World Development Indicators (WDI) database of the World Bank, while data on sector FDI and exports were collected from the Central Bank of Nigeria (CBN)

statistical bulletins. For the first hypothesis involving total FDI, the sample period is from 1980 to 2015. For the second hypothesis covering disaggregated FDI, the sample period is limited to 1981-2009. Both sample periods are dictated by the availability of data.

To avoid the risk of running regressions with nonstationary time series, we follow the literature (e.g., Bahmani-Oskooee and Hajilee, 2013) by testing for the stationarity of the variables. To this end, we apply the Ng and Perron (2001) and the augmented Dicky-Fuller (ADF) unit root tests.

Next, we employ the Autoregressive Distributed Lag (ARDL) cointegration approach (Pesaran and Shin, 1999; Pesaran *et al.*, 2001) to test for, and estimate, both the long- and short-run relationships between the variables. Unlike other cointegration methods, the ARDL approach is the most suitable for testing the long-run relationship among the variables when it is not known with certainty whether the regressors are purely I(0), purely I(1) or mutually cointegrated, as long as none of the regressors is integrated of order two (De Vita and Abbott, 2002). Fousekis *et al.* (2016) highlight several advantages of the ARDL approach to cointegration testing. It performs better in small samples compared to alternative multivariate cointegration procedures and is more efficient than the standard Engle and Granger two step approach. To illustrate, the ARDL (p, q) cointegration model with two time series y_t and x_t ($t = 1, 2, \dots, T$) has the following form:

$$\Delta y = \alpha_0 + \rho y_{t-1} + \theta x_{t-1} + \gamma z_t + \sum_{j=1}^{p-1} a_j \Delta y_{t-j} + \sum_{j=0}^{q-1} \pi_j \Delta x_{t-j} + e_t \quad (7)$$

where z_t is a vector of deterministic regressors, and e_t is a random disturbance term (an *iid* stochastic process).

Pesaran *et al.* (2001) show that the null of ‘no cointegration’, i.e., $H_0 : \rho = \theta = 0$, against the alternative hypothesis $H_1 : \rho \neq 0$ or $\theta \neq 0$, can be tested by employing a modified

F-test. Alternatively, the *t-BDM* test proposed by Banerjee *et al.* (1998), which tests the null of no cointegration $\rho = 0$ against the alternative $\rho < 0$, can be employed. The test procedure involves an upper bound and a lower bound. If the estimated value of the modified *F* or *t-BDM* statistic exceeds the upper critical bound then the null is rejected (i.e., y_t and x_t are cointegrated), if it lies below the lower critical bound the null cannot be rejected (i.e., y_t and x_t are not cointegrated), and if it lies between the critical bounds the test is inconclusive.

In terms of model selection, we apply the Akaike information criterion (AIC) rather than the Schwarz information criterion (SIC) since although by leading to lower order models for forecasting the latter has been found to be preferable judged on its ability to predict future values of the time series (see, e.g., Koehler and Murphree, 1988), if the chief objective is to explain the nature of the system generating the series – as in our case - the AIC is preferable given that SIC is stricter than AIC in penalising loss of degrees of freedom.

4. Results

4.1 Unit roots and ARDL cointegration

In Table I, the unit root test results for the variables in the first three models (equations 1-3) are presented while Table II summarises the unit root test results for the variables in equations 4 to 6. Despite a few discrepancies between the ‘constant only’ and ‘constant and trend’ specifications, the Ng-Perron and ADF unit root test results show that all the variables are $I(0)$ or $I(1)$ but not $I(2)$.

[Insert Table I and Table II here]

Having confirmed that none of the variables are integrated of order two (or higher), we proceed to apply the ARDL bounds test for cointegration.

Table III presents the estimated values of the F and t -BDM statistics of all the models at the 1%, 5% and 10% level of significance. For the models relating to the first hypothesis (that is, models 1 to 3), the test statistics exceed the upper critical bounds at the 10% level of significance in the total export and the oil export models. However, for the non-oil export model, the t -BDM and F statistics fall below the lower critical bound value at the customary significance levels. We, therefore, conclude that there exists no long-run relationship between non-oil exports and the independent variables in the model.

In the models relating to the second hypothesis (that is, models 4 to 6), the estimates of the F and t -BDM statistics show that there is a cointegrating relationship in the total export and the oil export models at the 1% significance level. For the non-oil export model, the two test statistics, again, fall below the lower critical bounds, hence we conclude that there is no long-run relationship between non-oil exports and the independent variables.

There are good explanations as to why there may be no cointegrating relationship between both aggregate and disaggregated FDI and non-oil exports in Nigeria. First, although FDI, especially FDI in the manufacturing sector, may be - at least in theory - expected to boost exports either through the exports of the MNE's affiliates or spillovers on the export capacity of domestic firms, the size of the investment and the motivation of the foreign investors matter in determining the impact of FDI on exports. If the level of FDI into the manufacturing sector is low, or the motivation for FDI is purely resource or market seeking or there are only few linkages between the domestic firms and manufacturing FDI, then it is possible that even manufacturing FDI may not have any long-run relationship with non-oil exports.

[Insert Table III here]

The diagnostic tests presented in Table IV suggest that there is no serial correlation and heteroscedasticity, and that normality holds. Furthermore, the plots of the cumulative sum (CUSUM) and the cumulative sum of squares (CUSUMQ) presented in Figure 4 show that there is no evidence of parameter instability.

[Insert Table IV and Figure 4 here]

4.2 Aggregate FDI and exports

Table IV presents the estimates of the long-run relationships and of the error correction models (ECMs) of our cointegrating models of aggregate FDI (we, therefore, do not estimate the non-cointegrating model of equation 3).

Looking first at Panel A of Table IV, in the first two models (1-2) pertaining to our first hypothesis (whether aggregate inward FDI has a positive effect on the volume of total exports and oil exports in Nigeria), we find that total FDI has a positive and highly statistically significant long-run relationship with both total exports as well as oil exports (not surprising given the prominent position of the oil sector in the Nigerian economy), with estimated coefficients of 1.150 and 1.209, respectively. This is consistent with our *a priori* expectation that the inflow of FDI will increase the volume of exports in the host economy. This result is supported by the similar findings of Wang *et al.* (2007) for China, and Leichenko *et al.* (1997) for the U.S.

We also find that the estimated REER coefficient in both model 1 and model 2, is not statistically significant, suggesting that in the case of Nigeria the exchange rate does not play a significant role in export performance. Contrary to our *a priori* expectations, our results also suggest that there exists a long-run negative relationship between GDP and both total exports and oil exports, with highly statistically significant coefficients of -3.679 and -3.828,

respectively. Dodaro (1993) argues that this may be because an increase in GDP could boost aggregate domestic demand, which in turn may make firms focus more on the domestic market and less on international trade. This explanation bears relevance in the context of Nigeria, which is a very large and highly populated country. Also, in their analysis covering over 90 countries, Anwar and Sampath (2001) found that the relationship between GDP and exports, while positive in many countries, is negative for less developed economies.

The associated ECM results are presented in Panel B of Table IV. The short-run coefficients suggest that $D(LNFDI)$ has a short-run positive impact on total exports and oil exports while $D(LNREER)$ and $D(LNGDP)$ do not have a significant impact on total or oil exports in the short-run. The Error Correction Term (ECT) of -0.569 and -0.543 for the total export and oil export models, respectively, are statistically significant and suggest that it takes just short of two years for the adjustment from short-run disequilibrium to long-run equilibrium to be achieved.

4.3 Sectoral FDI and exports

Our bound tests for cointegration suggest that there exists a long-run relationship between sectoral FDI and total exports (model 4) and sectoral FDI and oil exports (model 5), but we do not find evidence of a cointegrating relationship between sectoral FDI and non-oil exports (model 6). Hence, we only estimate model 4 and model 5.

The short-run coefficients display mixed effects, significant only for $D(LNMFDI)$ and $D(LNSFDI)$ at 1%, and $D(LNREER)$ at 10%. Such effects are, by their very nature, short-lived. The ECT in both model 4 and model 5 is negative and statistically significant, as to be expected, with estimated values suggesting that about 90% and 63% of the adjustment from

short-run disequilibrium to long-run equilibrium in the total export and oil export models, respectively, is completed within one year.

Our interest centres on the estimated long-run coefficients, which suggest that it is only primary sector FDI (*LNPFDI*) and manufacturing sector FDI (*LNMFDI*) that have a positive and statistically significant long-run impact on both total exports and oil exports. The effect is much more significant (at the 1% level) and more pronounced for *LNMFDI*, with an estimated coefficient of 1.486 and 1.526 in model 4 and 5, respectively. It is somewhat surprising that primary sector FDI has a positive long-run effect on total exports and oil exports only at the 10% significance level, with a coefficient of a smaller magnitude than manufacturing sector FDI given that the oil sector receives the greatest share of primary sector FDI, and it is well documented that FDI in primary sector in developing countries is usually export-oriented (see Hirschman, 1958; and Alfaro, 2003). However, the much higher and more significant elasticities for manufacturing sector FDI may be explained by the fact that manufacturing FDI helps in facilitating or constructing infrastructure that facilitates exports of all commodities as well as oil (Aitken *et al.*, 1997). This is probably the case from our data, as our results suggest that, in the case of Nigeria, manufacturing FDI has a strong positive effect on total exports in general, and oil exports in particular (though not for non-oil exports, which did not bear a cointegrating relationship).

The coefficient of service sector FDI (*LNSFDI*), on the other hand, suggests that there is no statistically significant long-run relationship between service sector FDI and total exports or oil exports (though the short-run coefficients are significant with a negative and positive effect in model 4 and 5, respectively). This result is not farfetched given that the service sector is still relatively underdeveloped in Nigeria. The estimated long-run REER coefficient (*LNREER*) is still negative but now statistically significant at the 10% level in the

two models (4 and 5), while the estimated long-run GDP coefficient (*LNGDP*) is statistically insignificant in the two models.

Taken together, our results suggest that while aggregate FDI has a positive long-run relationship with total exports in Nigeria, when disaggregated by export categories this relationship is significant only for oil exports. Moreover, when total FDI is disaggregated into sectors, the FDI-export nexus holds for primary and manufacturing FDI targeted at oil exports, not for non-oil exports. Services sector FDI does not bear any long-term relationship with Nigerian exports, whether oil or non-oil.

5. Conclusion

This study examines the direct impact of FDI on Nigeria's export performance by testing the effect of aggregate and sectoral FDI on total and disaggregated exports. By employing the ARDL cointegration technique, we find that aggregate FDI has a positive and statistically significant long-run relationship with total exports and oil exports. The evidence further suggests that there is no statistically significant impact of FDI on non-oil exports. A similar result is found when FDI is disaggregated according to sector. Both primary sector FDI and manufacturing sector FDI have a positive long-run relationship with total exports and oil exports. Service sector FDI, on the other hand, appears not to have a statistically significant long-run impact on exports at either aggregate or disaggregated level.

These findings lend themselves to straightforward policy recommendations. The Nigerian government ought to create incentives to attract the type of FDI that has a greater impact on exports, namely primary sector FDI and especially manufacturing sector FDI. That said, in order to reduce the dependence on the oil sector, it is important that Nigerian policymakers interested in export diversification, specifically incentivise the production and

exports of non-oil export goods. Investment promotion priorities should be given to foreign firms who are more likely to increase local industrialisation. In addition, efforts should be made to improve the absorptive capacity of Nigeria and domestically-owned firms, so as to increase the probability that domestically-owned firms benefit from spillovers arising from the presence of these foreign firms.

Data availability permitting, investigation of the indirect spillover effect of FDI on domestic (Nigerian) firms' exports, would constitute a valuable extension of our analysis. Also, our study did not consider any possible nonlinearity in the relationship in question. It follows that investigation of a potentially significant nonlinear relationship between FDI and host country exports, at both aggregate and disaggregated level, provides a profitable avenue for future research.

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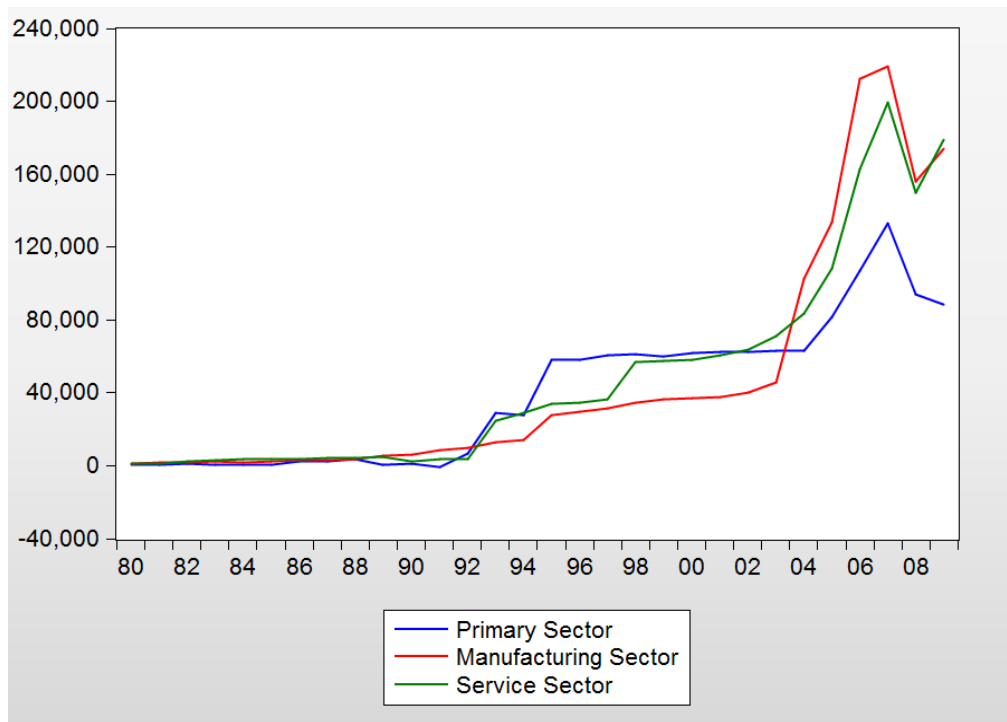
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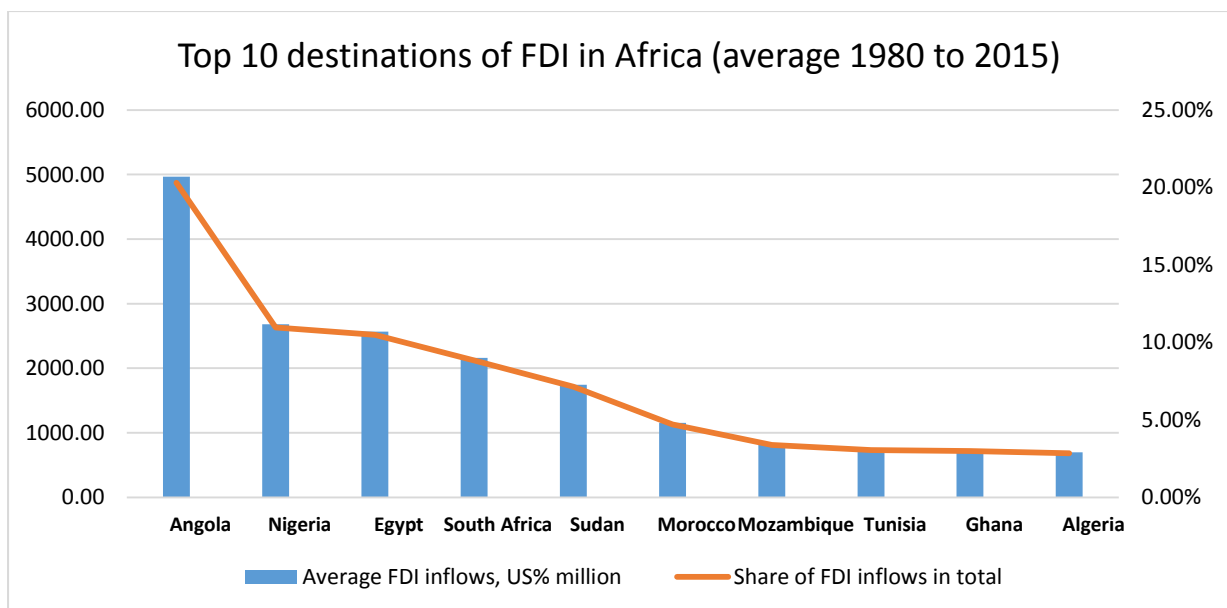
Figure 1

Inward FDI flows in Nigeria by sector from 1980 to 2009.



Source: CBN Statistical Bulletin 2009.

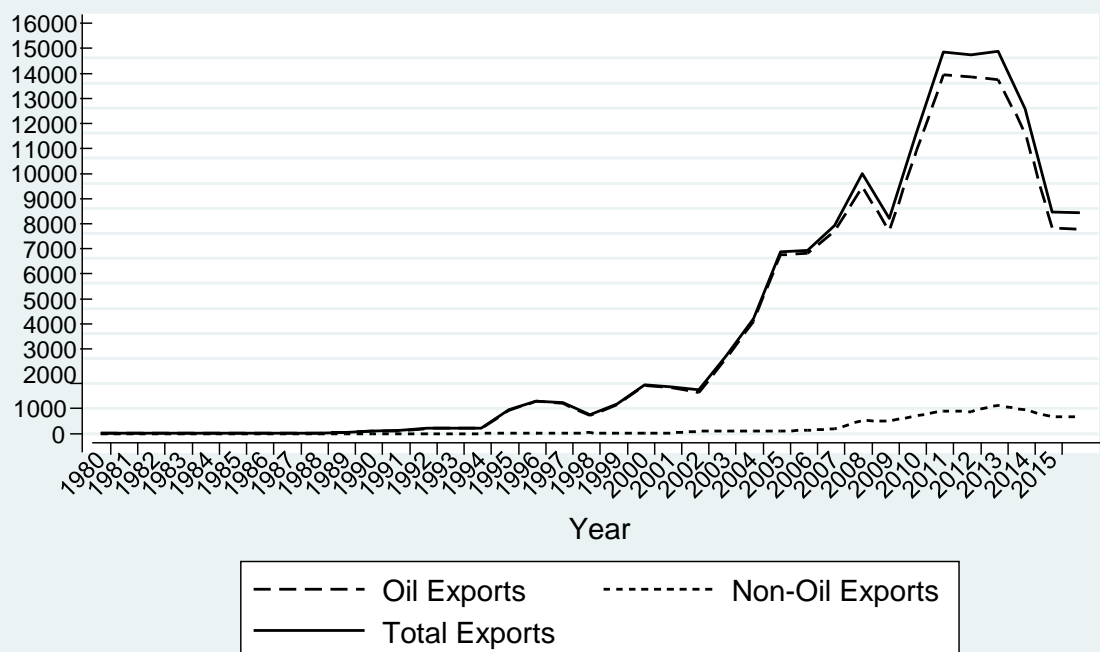
Figure 2



Source: Calculations based on data from UNCTAD.

Figure 3

Summary of export in Nigeria: Total, Oil and Non-Oil (N' Billion)

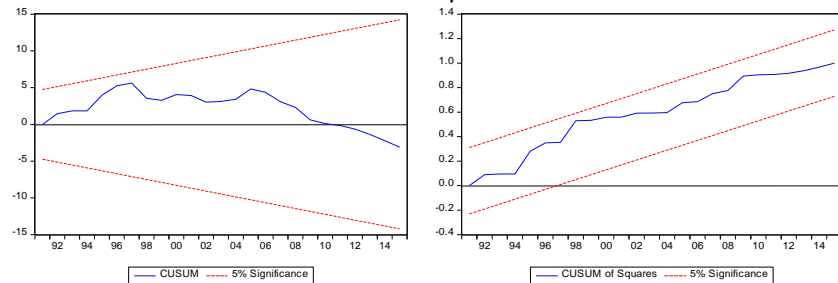


Source: CBN Statistical Bulletin 2016.

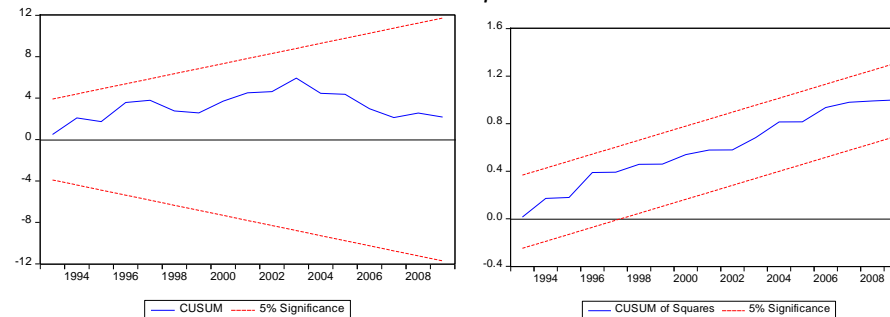
Figure 4

CUSUM and CUSUMSQ for the ARDL Models

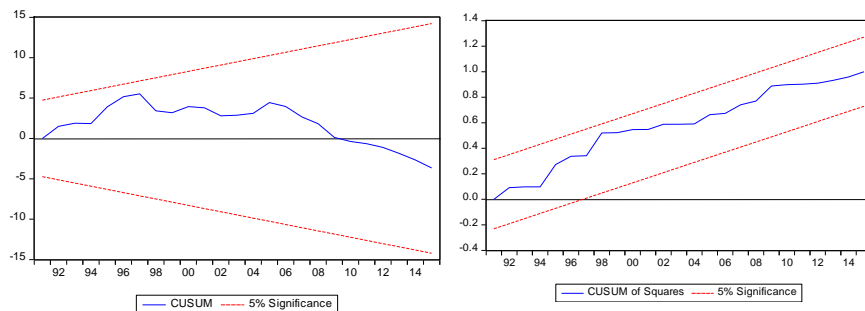
Aggregate FDI
Total export



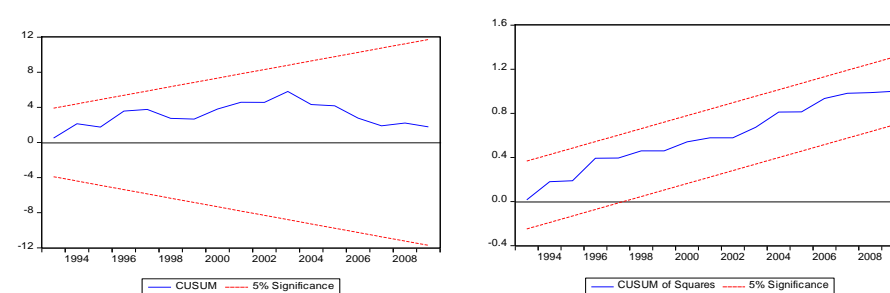
Sectoral FDI
Total export



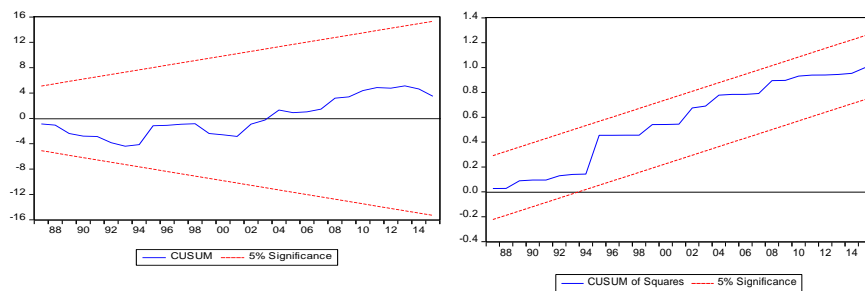
Oil export



Oil export



Non-oil export



Non-oil export

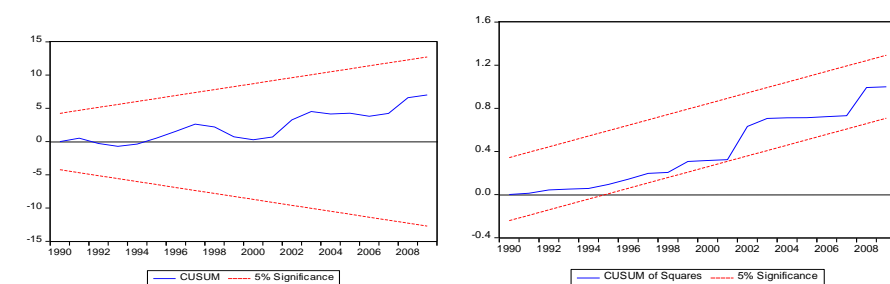


Table I.

Ng and Perron (2001) and ADF unit root tests.

Variables	constant only						constant and time trend					
	MZ_a^{GLS}	MZ_t^{GLS}	MSB^{GLS}	MP_T^{GLS}	ADF	k	MZ_a^{GLS}	MZ_t^{GLS}	MSB^{GLS}	MP_T^{GLS}	ADF	k
<i>LNFDI</i>	-0.971	-0.597	0.615	20.348	-0.814	0	-13.455	-2.526	0.187	7.1473	-7.241***	1
<i>LNOEXP</i>	0.438	0.458	1.046	66.978	-0.942	0	-5.144	-1.357	0.263	16.687	-1.233	0
<i>LNNEXP</i>	0.677	0.628	0.927	56.876	-0.542	0	-11.136	-2.260	0.203	8.668	-3.426*	0
<i>LNTEXP</i>	0.500	0.531	1.061	69.657	-0.898	0	-5.329	-1.401	0.262	16.310	-1.330	0
<i>LNREER</i>	-3.977	-1.396	0.351	6.171	-1.930	0	-5.809	-1.653	0.284	15.592	-1.771	0
<i>LNGDP</i>	1.096	0.764	0.697	37.897	1.676	0	-1.365	-0.689	0.504	50.729	-0.205	0
Δ <i>LNFDI</i>	-13.475**	-2.583***	0.191	1.864**	-10.791***	1	-11.699	-2.390	0.204	7.936	-10.610***	1
Δ <i>LNOEXP</i>	-16.460***	-2.839***	0.172***	1.598***	-5.789***	1	-16.598*	-2.826*	0.170*	5.808*	-4.503***	1
Δ <i>LNNEXP</i>	-6.277*	-1.757*	0.279	3.948*	-6.468***	1	-16.240*	-2.786*	0.171*	5.980*	-6.355***	1
Δ <i>LNTEXP</i>	-16.432***	-2.845***	0.173***	1.568***	-5.849***	1	-16.585*	-2.834*	0.170*	5.760*	-4.547***	1
Δ <i>LNREER</i>	-15.825***	-2.812***	0.177**	1.549***	-4.403***	1	-15.929*	-2.822*	0.177*	5.721*	-4.416***	1
Δ <i>LNGDP</i>	-6.210*	-1.740*	0.280	4.014*	-4.824***	1	-15.651*	-2.743*	0.175*	6.135*	-5.330***	1
<i>Critical values</i>												
1%	-13.80	-2.58	0.174	1.78			-23.80	-3.42	0.143	4.03		
5%	-8.10	-1.98	0.233	3.17			-17.30	-2.91	0.168	5.48		
10%	-5.70	-1.62	0.275	4.45			-14.20	-2.62	0.185	6.67		

Notes: Δ denotes the first difference operator while k denotes the order of integration of the series. The critical values are from Ng and Perron (2001). ***, ** and * denote the rejection of the null of a unit root at the 1%, 5% and 10% significance level, respectively. Ng and Perron (2001) combine GLS detrending with SD to design a new test. The proposed test consists of a suite of four tests, namely MZ_a^{GLS} , MZ_t^{GLS} , MSB^{GLS} and MP_T^{GLS} to correct for size distortions when residuals are negatively correlated. For the ADF tests, a maximum lag length of eight lags was used.

Table II.

Ng-Perron (2001) and ADF unit root tests.

Variables	constant only						constant and time trend					
	MZ_a^{GLS}	MZ_t^{GLS}	MSB^{GLS}	MP_T^{GLS}	ADF	k	MZ_a^{GLS}	MZ_t^{GLS}	MSB^{GLS}	MP_T^{GLS}	ADF	k
<i>LNOEXP</i>	0.702	0.666	0.948	59.365	59.365	0	-8.540	-1.968	0.230	10.976	-2.527	0
<i>LNEXP</i>	1.207	0.956	0.791	47.698	47.698	0	-11.592	-2.404	0.207	7.875	-4.074	0
<i>LNTEXP</i>	0.737	0.705	0.957	60.825	60.825	0	-8.625	-1.983	0.229	10.867	-2.556	0
<i>LNREEF</i>	-3.381	-1.251	0.370	7.2105	7.2105	0	-5.369	-1.615	0.300	16.887	-1.755	0
<i>LNGDP</i>	0.893	0.440	0.492	21.593	21.593	0	-1.242	-0.518	0.417	39.704	-1.650	0
<i>LNPFDI</i>	-1.037	-0.542	0.523	16.343	16.343	0	-8.019	-1.934	0.241	11.537	-2.150	0
<i>LNMFDI</i>	0.553	0.383	0.692	33.867	33.867	0	-327.563***	-12.792***	0.0390***	0.290***	-3.754**	1
<i>LNSFDI</i>	0.969	0.689	0.711	38.180	38.183	0	-6.1605	-1.753	0.284	14.790	-1.909	0
Δ <i>LNOEXP</i>	-13.212**	-2.556**	0.193**	1.906**	1.906**	1	-12.817	-2.489	0.194	7.334	-4.574***	1
Δ <i>LNEXP</i>	-13.120**	-2.559**	0.195**	1.874**	1.874**	1	-12.353	-2.457	0.198	7.524	-6.494***	1
Δ <i>LNTEXP</i>	-13.192**	-2.555**	0.193**	1.907**	1.907**	1	-12.789	-2.486	0.194	7.353	-4.579***	1
Δ <i>LNREER</i>	-12.689**	-2.518**	0.198**	1.933**	1.933**	1	-12.727	-2.521	0.198	7.168	-3.880**	1
Δ <i>LNGDP</i>	-9.667**	-2.158**	0.223**	2.686**	2.686**	1	-12.840	-2.510	0.195	7.222	-5.163***	1
Δ <i>LNPFDI</i>	-13.037**	-2.544**	0.195**	1.911**	1.911**	1	-12.998	-2.548	0.196	7.0136	-5.955***	1
Δ <i>LNMFDI</i>	-13.359**	-2.583***	0.193**	1.839**	1.839**	1	-13.366	-2.582	0.193	6.831	-4.462***	1
Δ <i>LNSFDI</i>	-13.037**	-2.541**	0.194**	1.922**	1.922**	1	-12.900	-2.539	0.196	7.064	-5.769***	1
<i>Critical values</i>												
1%	-13.80	-2.58	0.174	1.78			-23.80	-3.42	0.143	4.03		
5%	-8.10	-1.98	0.233	3.17			-17.30	-2.91	0.168	5.48		
10%	-5.70	-1.62	0.275	4.45			-14.20	-2.62	0.185	6.67		

Notes: Δ denotes the first difference operator while k denotes the order of integration of the series. The critical values are from Ng and Perron (2001). ***, ** and * denote the rejection of the null hypothesis of a unit root at the 1%, 5% and 10% significance level, respectively. Ng and Perron (2001) combine GLS detrending with SD to design a new test. The proposed test consists of a suite of four tests, namely MZ_a^{GLS} , MZ_t^{GLS} , MSB^{GLS} and MP_T^{GLS} to correct for size distortions when residuals are negatively correlated. For the ADF tests, a maximum lag length of eight lags was used.

Table III.

Bounds testing for cointegration (equations 1-6).

<i>Aggregate FDI</i>							<i>Sectoral FDI</i>					
<i>Total Export (1)</i>		<i>Oil Export (2)</i>		<i>Non-Oil Export(3)</i>			<i>Total Export (4)</i>		<i>Oil Export (5)</i>		<i>Non-Oil Export(6)</i>	
F-statistic	4.209		4.060		1.110		7.081		7.148		1.122	
Critical Value Bounds							Critical Value Bounds					
Significance	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound
10%	2.97	3.74	2.97	3.74	2.97	3.74	2.58	3.86	2.58	3.86	2.58	3.86
5%	3.38	4.23	3.38	4.23	3.38	4.23	3.12	4.61	3.12	4.61	3.12	4.61
2.5%	3.8	4.68	3.8	4.68	3.8	4.68	4.54	6.37	4.54	6.37	4.54	6.37
1%	4.3	5.23	4.3	5.23	4.3	5.23	2.58	3.86	2.58	3.86	2.58	3.86
t-BDM	-3.911		-3.924		-1.090		-5.926		-5.914		-2.477	
Critical Value Bounds							Critical Value Bounds					
Significance	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound	I(0) Bound	I(1) Bound
10%	-3.13	-3.84	-3.13	-3.84	-3.13	-3.84	-2.57	-3.86	-2.57	-3.86	-2.57	-3.86
5%	-3.41	-4.16	-3.41	-4.16	-3.41	-4.16	-2.86	-4.19	-2.86	-4.19	-2.86	-4.19
2.5%	-3.65	-4.42	-3.65	-4.42	-3.65	-4.42	-3.13	-4.46	-3.13	-4.46	-3.13	-4.46
1%	-3.96	-4.73	-3.96	-4.73	-3.96	-4.73	-3.43	-4.79	-3.43	-4.79	-3.43	-4.79

Table IV.

Error correction and cointegration models.

<i>Panel A: Long-run coefficients</i>					
<i>Aggregate FDI</i>			<i>Sectoral FDI</i>		
Variable	Total Export (1)	Oil Export (2)	Variable	Total Export (4)	Oil Export (5)
LNREER	-0.098 (0.681)	-0.117 (0.652)	LNREER	-0.303* (0.060)	-0.292* (0.073)
LNGDP	-3.679*** (0.000)	-3.828*** (0.000)	LNGDP	-0.5693 (0.544)	-0.648 (0.497)
LNFDI	1.150*** (0.001)	1.209*** (0.002)	LNPFDI	0.268* (0.070)	0.258* (0.085)
C	41.637*** (0.000)	41.211*** (0.000)	LNMFDI	1.486*** (0.000)	1.526*** (0.000)
			LNSFDI	-0.300 (0.242)	-0.306 (0.240)
			C	7.785 (0.711)	9.493 (0.657)
<i>Panel B: Short-run coefficients</i>					
<i>Aggregate FDI</i>			<i>Sectoral FDI</i>		
Variable	Total Export (1)	Oil Export (2)	Variable	Total Export (4)	Oil Export (5)
D(LNREER)	-0.129 (0.431)	-0.144 (0.392)	D(LNREER)	-0.226* (0.074)	-0.217* (0.091)
D(LNGDP)	0.008 (0.992)	0.073 (0.938)	D(LNGDP)	-0.503 (0.463)	-0.521 (0.454)
D(LNFDI)	0.251** (0.048)	0.254* (0.051)	D(LNPFDI)	-0.012 (0.851)	-0.011 (0.864)
ECT(-1)	-0.569*** (0.000)	-0.543*** (0.000)	D(LNMFDI)	1.142*** (0.000)	-0.230*** (0.005)
			D(LNSFDI)	-0.624*** (0.003)	1.164*** (0.000)

			ECT(-1)	-0.903*** (0.000)	-0.633** (0.004)
<i>Diagnostics</i>					
SC	0.299 (0.589)	0.337 (0.566)	SC	2.179 (0.147)	2.1834 (0.147)
HETER	1.377 (0.254)	1.261 (0.306)	HETER	1.017 (0.464)	1.033 (0.454)
NORM	1.162 (0.559)	0.953 (0.620)	NORM	0.352 (0.838)	0.298 (0.861)

Notes: Probabilities values are presented in parenthesis. ***, ** and * denote the rejection of the null hypothesis of a unit root at the 1%, 5% and 10% significance level, respectively. The optimal lag structure is selected by AIC, starting with max 2 lags. SC denotes the Breusch and Godfrey serial correlation test, HETER denotes the Breusch and Pagan heteroscedasticity test, and NORM denotes the Jarque–Bera test for normality.