

# Birthweight, HIV exposure and infant feeding as predictors of malnutrition in Botswanan infants

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1 **Title Page**

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3 **Title**

4 Predictors of malnutrition in Botswanan infants during their first 1000 days; the contribution of  
5 birthweight, infant feeding practices and HIV-exposure

6

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20 **Authors' Roles:**

21 PC, SLE, JAS and DM were responsible for the study design and protocol and tools used in the  
22 study. PC was responsible for data collection. CE was responsible for nutritional analysis of the  
23 cereal samples consumed by study participants. Data analysis and review was conducted by PC,  
24 JAS and SLE. All authors were responsible for completing the manuscript.

25 **Key words:** child undernutrition, malnutrition, HIV, infant feeding practices, 1000 days, Botswana

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## 31 **Abstract**

32 **Background:** A better understanding of the nutritional status of infants who are HIV-Exposed-  
33 Uninfected (HEU) and HIV-Unexposed-Uninfected (HUU) during their first 1000 days is a key to  
34 improving population health, particularly in sub-Saharan Africa.

35 **Methods:** A cross-sectional study compared nutritional status, feeding practices and determinants of  
36 nutritional status of HEU and HUU infants residing in representative selected districts in Botswana  
37 during their first 1000 days of life. Four hundred and thirteen infants (37.3% HIV-exposed), aged 6-  
38 24 months attending routine child health clinics were recruited. Anthropometric, 24-hour dietary  
39 intake and socio-demographic data was collected. Anthropometric z-scores were calculated using  
40 2006 WHO growth standards. Modelling of the determinants of malnutrition was undertaken using  
41 logistic regression.

42 **Results:** Overall, prevalence of stunting, wasting and underweight were 10.4%, 11.9% and 10.2%  
43 respectively. HEU infants were more likely to be underweight (15.6% vs. 6.9%), ( $p < 0.01$ ) and stunted  
44 (15.6% vs. 7.3%), ( $p < 0.05$ ) but not wasted ( $p = 0.14$ ) than HUU infants. HEU infants tended to be  
45 formula fed (89.4%) whereas HUU infants tended to breastfeed (89.6%) for the first six months  
46 ( $p < 0.001$ ). Significant predictors of nutritional status were HIV exposure, birthweight, birth length,  
47 Apgar score and mother/caregiver's education with little influence of socioeconomic status.

48 **Conclusions:** HEU infants aged 6-24 months had worse nutritional status compared to HUU infants.  
49 Low birthweight was the main predictor of undernutrition in this population. Optimisation of infants'  
50 nutritional status should focus on improving birthweight. In addition, specific interventions should  
51 target HEU infants in order to eliminate growth disparity between HEU and HUU infants.

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## 60 **Introduction**

61 Globally, under-five mortality declined from 90 to 43 deaths per 1000 live births between 1990 and  
62 2015 <sup>(1)</sup>. However, in sub-Saharan Africa under-five mortality still remains higher at 86 deaths per  
63 1000 live births <sup>(1)</sup>. Mortality in children aged less than five years is mainly attributed to  
64 undernutrition, with 45% of these deaths being preventable through optimal nutrition, especially in  
65 the first 1000 days (the period from conception to the child's second birthday) <sup>(2, 3-4)</sup>.

66 Human Immunodeficiency Virus/Acquired Immune Deficiency Syndrome (HIV/AIDS) is still a  
67 major health challenge in Botswana <sup>(5-6)</sup>. Strategies including, prevention of mother-to-child  
68 transmission (PMTCT) of HIV have been highly successful in Botswana, reducing mother-to-child  
69 transmission rates to approximately 2.6% <sup>(7-8)</sup>. Without PMTCT strategies, HIV transmission from  
70 mother to child could be as high as 25% <sup>(8)</sup>. However, this success has resulted in the increase in the  
71 population of HIV-exposed but uninfected (HEU) infants. <sup>(9-10)</sup>. Health and/or nutritional issues  
72 unique to HEU infants will have major population health implications as their numbers increase <sup>(11-</sup>  
73 <sup>12)</sup>. Currently, the health and nutritional consequences of HIV-exposure are largely under study <sup>(10, 12)</sup>.  
74 However, a higher risk for mortality in HEU compared to HIV-unexposed uninfected (HUU) infants  
75 has been previously reported <sup>(13-16)</sup>. Risk of mortality can be modified by optimising nutritional status  
76 of infants, this requires a good understanding of context specific patterns and determinants of  
77 undernutrition in this group <sup>(17)</sup>.

78 Studies conducted in other African countries comparing the nutritional status of HEU and HUU  
79 infants show large variations in the levels of undernutrition <sup>(12, 18-20)</sup>. Majority of these studies were  
80 conducted before ART was widely available to mothers and infants <sup>(12, 18-20)</sup>. In contrast, ART is  
81 available to approximately 92% of pregnant women in Botswana <sup>(8)</sup>. Monitoring and management of  
82 infant health and nutrition is intensive and widely accessible <sup>(21)</sup>. The same conditions are often not  
83 present in other sub-Saharan African countries with high HIV prevalence. However, the level of  
84 mortality in HEU infants in Botswana is comparable or higher than in other sub-Saharan African  
85 countries <sup>(13)</sup>. Furthermore, the feeding policy adopted for HEU infants in Botswana is unique, as it  
86 has inadvertently undermined breastfeeding levels through provision of free formula <sup>(22)</sup>. Currently,  
87 the nutritional status of HEU and HUU in Botswana have not been well documented. Therefore,  
88 understanding nutritional status and its determinants between HEU and HUU infants in Botswana is  
89 important for informing policies and interventions which can be used to achieve comparable growth  
90 between these infants if such differences exist. Thus helping reduce the risk of mortality in HEU  
91 infants. This study sought to investigate the patterns of undernutrition per HIV-exposure within

92 context of feeding practices in infants aged 6- 24 months in selected districts in Botswana. In  
93 addition, this study, also aimed to identify determinants of nutritional status in these infants.

## 94 **Methods**

### 95 **Study participants and population**

96 The study was conducted in Botswana using a comparative cross-sectional study design between  
97 December 2014 and February 2015 in 19 different government health facilities of varying sizes  
98 (hospital, primary hospital, clinics and/or health posts) located across the four districts (Kweneng-  
99 East, Kgatleng, Selebi Phikwe and Francistown). Health facilities in districts with high HIV  
100 prevalence in the adult population were selected in order to obtain an adequate number of HEU  
101 infants. Prevalence of HIV in these districts ranged from 26.3% in Kweneng East to 39.6% in Selebi  
102 Phikwe They were selected as having higher HIV prevalence than the national average, in order to  
103 ensure an appropriate sample of HEU infants. These four districts were selected to represent urban,  
104 semi-urban and rural areas. Kweneng-East is mainly rural with some semi-urban locations. Kgatleng  
105 is mainly rural, Selebi Phikwe is semi-urban while Francistown is mostly urban. These locations span  
106 the eastern hardveld where at least 80% of the population of Botswana live <sup>(23-24)</sup>. All caregivers from  
107 the general population with infants aged 6-24 months attending their monthly growth monitoring in  
108 a health facility, were invited to participate in the study. Eligible caregivers had to be citizens of  
109 Botswana, aged over 18 years and were the infant's parent and/or legal guardian. There were no other  
110 exclusion criteria. The participants were approached as they arrived at the health facility. Children in  
111 Botswana, aged 0-59 months attend routine monthly growth monitoring in government health  
112 facilities across the country. When more participants than required showed interest in the study,  
113 simple randomisation was used to select participants by allocating each participant a number.

### 114 **Sample size**

115 A representative sample of infants in selected districts was stratified according to the population of  
116 the infants aged under five years in each district based upon data supplied by the Ministry of Health  
117 and Wellness in Botswana (Nutrition and Food Control division). Therefore, a district with a higher  
118 number of under-fives had a larger representation within the sample. In addition, the composition of  
119 the sample within each district was selected such that it represented the proportions of infants  
120 attending each type of health facility (hospital, primary hospital, clinics and/or health posts) within  
121 that district. Therefore, a type of health facility receiving a higher number of infants would have a  
122 higher share of the sample within each district.

123

124 To facilitate a logistic regression analysis, an adequate sample size assuming a medium size

125 relationship between the dependant variables (underweight, stunting and wasting) and independent  
126 variables and,  $\alpha = 0.05$  and  $\beta = 0.20$  was taken to be  $N \geq 50 + 8m$  (where  $m$  is the number of  
127 independent variables) <sup>(25)</sup>. In total, 44 potential independent variables were identified *a priori* to the  
128 data collection; resulting in a minimum sample size of 402 caregiver-infant pairs (see Table S1). In  
129 addition, oversampling by 10% was also employed to counter missing data. Independent variables  
130 identified *a priori* and known to affect undernutrition in infants such as birthweight, sex, and maternal  
131 age, care giver education level and socio economic factors were included <sup>(26-28)</sup>. These variables were  
132 derived from data collection (anthropometry, dietary recall, interview of caregivers) and review of  
133 the child health card. However, due to the cross-sectional nature of the study, maternal nutrition and  
134 health variables prior to the study, such as during pregnancy were not available. HIV-exposure was  
135 maintained in all analysis as it was a variable of interest.

### 136 **Procedures**

137 Participants were recruited during their infant's free monthly routine health check-up at a health  
138 facility. In total 419 participants were approached to take part in the study. Five infants with an  
139 undocumented HIV status and/or missing PCR DNA/rapid HIV tests were not enrolled into the study.  
140 Of all the participants approached, only one declined to take part in the study. The final sample size  
141 was, therefore, 413 infants.

142 Data were collected by the lead author and two trained assistants using a structured interview with  
143 the caregiver and review of each child's health card. All caregivers in Botswana are given and keep  
144 a health card for their infant at birth. This card contains details such as birthweight and length,  
145 vaccinations, monthly weight and feeding practices. HIV-exposure was determined from the child's  
146 health card as per the latest DNA/PCR or rapid test result. HIV negative mothers were tested every  
147 three months for HIV during antenatal care, with the latest test at 36 weeks documented in the child's  
148 health card. Socio-demographic characteristics, feeding practices and health history as potential  
149 independent variables were collected from the caregiver and the health card. Anthropometric  
150 measures of length/height and weight were measured in duplicate from all the infants as per WHO  
151 standard procedure <sup>(29)</sup> using standardised equipment. Weight was measured to the nearest 0.05g using  
152 calibrated Seca<sup>®</sup> Scales 385 and 875 (Seca gmbh & co, Hamburg, Germany) and length/height was  
153 measured to the nearest 1 mm using Seca<sup>®</sup> measuring board 417 (Seca gmbh & co, Hamburg,  
154 Germany) and Seca<sup>®</sup> stadiometre, Seca 217 (Seca gmbh & co, Hamburg, Germany). Length for age  
155  $z$ -scores (*LAZ*), weight for age  $z$ -scores (*WAZ*) and weight for length  $z$ -scores (*WLZ*) were calculated  
156 according to the 2006 WHO child growth standards using the WHO Anthro 2005 programme, Beta  
157 version <sup>(30)</sup>. Stunting, underweight and wasting was determined at  $z$  score  $< -2$  SD based on *LAZ*,  
158 *WAZ* and *WLZ* respectively.

159 A modified USDA five step multiple Pass 24-hour dietary recall protocol <sup>(31)</sup> was used to measure  
160 infant's current nutritional intake as recalled by the caregiver. A similar multiple pass 24-hour dietary  
161 recall was validated in Ugandan children and was found to be valid in assessing dietary intake of  
162 infants residing in communities with similar diets <sup>(32)</sup>. Dietary diversity was calculated by allocating  
163 a score for consumption of food from one of the seven food groups (Grains, roots and tubers: Legumes  
164 & nuts: Dairy products: Flesh foods: Eggs: vitamin A rich fruits and vegetables: other fruits and  
165 vegetables) in the preceding 24 hours <sup>(33)</sup>. Therefore, resulting in a maximum possible score of 7, an  
166 infant's diet scoring 4 or more is considered diverse <sup>(33)</sup>. In addition, to dietary diversity <sup>(33)</sup>, Nutritics<sup>®</sup>  
167 software <sup>(34)</sup>, was used to derive the energy and protein intake of each infant. Nutritional information  
168 of foods consumed was derived from packaging, data from South African Composition Database <sup>(35)</sup>  
169 and McCance and Widdowson's composition of foods databases <sup>(36)</sup>. Cereals such as sorghum and  
170 fortified sorghum were consumed by majority of infants but nutritional content was not available.  
171 Therefore, cooked samples of these were weighed, frozen then freeze dried and analysed in the  
172 laboratory for protein per 100 grams using the Flash EA1112 nitrogen elemental analyser (Soeks FL  
173 33334, USA). Energy per 100 grams was analysed using Parr 6300 Oxygen bomb calorimetre (Parr  
174 Instrument Co., Moline, Illinois, USA).

175 Data was entered into SPSS version 22 software <sup>(37)</sup> for analysis and 10% of this data was randomly  
176 selected using a computer number generator and then screened for accuracy by the co-authors.

## 177 **Ethics**

178 Ethical approval was received both from the University of Nottingham's Medical School Research  
179 Ethics Committee and the Health Research and Development Committee in Botswana. Informed  
180 consent was obtained from all caregivers. The two assistants were trained in seeking informed  
181 consent. When inappropriate feeding and/or malnutrition were identified the caregiver was briefly  
182 counselled by the lead author, who is also a registered dietitian. The caregiver was then referred to  
183 the health facility for further follow-up and this was documented in the child's health card to ensure  
184 continuity of care.

## 185 **Statistical methods**

186 Data was analysed using Statistical Package for Social Sciences, SPSS version 22 <sup>(37)</sup>. A case-control  
187 analysis approach was employed where HEU and HUU infants were compared for outcomes of  
188 interest. Baseline data is described as per HIV exposure. Chi square, was used to test for proportions  
189 between the two groups (HEU and HUU infants) to determine prevalence of underweight, wasting  
190 and stunting. Continuous variables were analysed using Kolgorov-Smirnov test to determine whether  
191 the distribution was Gaussian or not. Independent samples t-test or Mann-Whitney U test were used  
192 to test for differences between the two groups for parametric and non-parametric variables

193 respectively. Variation of the mean was presented as standard deviation. Forward logistic regression  
194 was performed to determine predictors of stunting, underweight and wasting. The threshold for  
195 introducing the variables into the logistic regression model was set at  $p < 0.1$ . Cases with missing  
196 values for some of the independent variables were excluded. On this basis 86.2% of cases with no  
197 missing values were included in the analysis for each of the three dependent variables (stunting,  
198 underweight and wasting). Variables with missing data included feeding method at  $< 6$  months  
199 (2.6%), feeding method at 6-12 months (6.1 %), birthweight (4.1%), Apgar score (2.9%), and age at  
200 which complementary feeds were introduced (2.4%). One of the co-authors (JAS) had the overall  
201 oversight of the statistical methods and analysis. Statistical significance was taken at  $p < 0.05$  in all  
202 analysis.

## 203 **Results**

### 204 **Characteristics of participants**

205 A total of 413 participants were recruited, of which 154 were HEU (37.3%) and 259 were HUU  
206 (62.7%). Table 1 shows the characteristics of participants by HIV exposure. No significant  
207 differences were found between HEU and HUU infants in terms of age, proportions of sex,  
208 birthweight or length nor birthweight classification. However, HEU infants had significantly more  
209 siblings compared to HUU infants ( $p < 0.001$ ). In addition, HEU infants were more likely to have  
210 had a sibling who died compared to HUU infants ( $p < 0.05$ ).

211 As shown in Table 1, HIV positive mothers tended to be older at the time of the infant's birth  
212 ( $p < 0.001$ ). In addition, the primary caregivers of HEU infants had significantly lower education levels  
213 ( $p < 0.001$ ). No significant differences were found in other mother/caregiver and household  
214 characteristics between the two groups.

### 215 **Feeding practices**

216 Table 2 shows feeding practices of infants per HIV-exposure from birth to age at time of data  
217 collection. These feeding practices were self-reported by the caregiver and corroborated using data  
218 from each child's health card, where possible. HEU infants were more likely to be formula fed from  
219 birth and at 6-12 months compared to HUU infants ( $p < 0.001$ ). The remainder of the infants ( $n=11$ )  
220 not breastfeeding or formula feeding in the first twelve months were taking cow's milk. Of those  
221 infants aged more than 12 months, it was found that HUU infants were more likely to be breastfed  
222 compared to their HEU counterparts ( $p < 0.001$ ). Overall the energy and protein intake for male and  
223 female HEU and HUU infants were higher than recommended nutrient intakes (RNI for infants aged  
224 1-3 years). Average energy and protein intake was found to be higher in HEU compared to HUU  
225 infants for females and *vice versa* for males. However, both these differences did not reach statistical



226 significance. In addition, there were no significant differences between HEU and HUU infants in age  
227 at which the infant was introduced to complementary feeds. Dietary diversity was low for all infants,  
228 and there was no significant difference between HEU and HUU infants.

### 229 **Nutritional outcomes**

230 The prevalence of underweight was higher in HEU infants (Table 3;  $p < 0.01$ ). In addition, HEU infants  
231 also had significantly higher prevalence of stunting compared to HUU infants (15.6% vs. 7.3 %,  $p < 0.05$ ).  
232 Wasting prevalence was higher in HEU infants; however this did not reach statistical  
233 significance ( $p = 0.14$ ).

### 234 **Determinants of nutritional status**

235 The results of logistic regression to identify the determinants of underweight, stunting and wasting  
236 are shown in Tables 4, 5 and 6. Table 4 shows the determinants of underweight. The analysis revealed  
237 that infants living in homes where a child had previously died were over three times more likely to  
238 be underweight (adjusted OR 3.205, 95% CI 1.097- 9.362). However, a higher birthweight or birth  
239 length was negatively associated with underweight ( $p < 0.001$ ,  $p = 0.03$  respectively). Each kilogram  
240 higher weight reduced risk of underweight by 82% (OR 0.182, 95% CI 0.073 -0.450). Similarly, a  
241 1cm increase in birth length reduced risk by 10% (OR 0.899, 95% CI 0.818 -0.988). Importantly,  
242 HIV exposure, infant nutrient intakes, maternal and household factors were not associated with risk  
243 of underweight. Predictors for stunting as shown in Table 5, were consistent with the simple chi  
244 square analysis of prevalence. HEU infants were found to be more than twice as likely to be stunted  
245 compared to HUU infants (adjusted OR 2.361, 95% CI 1.105 -5.046). In addition, a lower level of  
246 mother/caregiver's education, and lower birthweight was associated with stunting. Again, nutrient  
247 intakes and other maternal and household factors were not significantly associated with risk of  
248 stunting. Wasting was more likely in infants with a high Apgar score, however residing in Kweneng  
249 East district (rural/semi urban) and having a higher birthweight was negatively associated with  
250 wasting. Each kilogram extra weight at birth reduced risk of wasting by 58% (adjusted OR 0.423,  
251 95%CI 0.205-0.872). HIV exposure, infant nutrient intake and other household and maternal factors  
252 were not significantly associated with risk of wasting.

253

### 254 **Discussion**

255 Our study has demonstrated that HEU infants aged 6-24 months have poor nutritional outcomes  
256 compared to HUU infants. This has implications for policy and programming because currently  
257 prevention of mother-to-child transmission of HIV in HEU infants is prioritised over achieving  
258 optimal nutritional status. This has inadvertently resulted in inequitable growth between HEU and

259 HUU infants. Data from 154 HEU infants and 259 HUU infants living in selected districts in  
260 Botswana demonstrated that HEU infants had higher prevalence of underweight and stunting. HEU  
261 infants were also more likely to formula feed in their first 12 months of life whereas HUU infants  
262 were more likely to breastfeed. Low birthweight was the strongest predictor of undernutrition in  
263 addition to HIV exposure, birth length, mother/care giver's education level, high Apgar score and  
264 residing in Kweneng East.

265 Prevalence of undernutrition in this study was higher in HEU infants compared to HUU infants during  
266 their first 1000 days. This is consistent with findings from a number of studies conducted in Zambia,  
267 Kenya, South-Africa, Uganda and Tanzania which have demonstrated that HEU infants have poor  
268 growth compared to HUU infants<sup>(9, 12, 20, 38-39)</sup>. A study in Kenyan infants found that HEU infants had  
269 poor nutritional outcomes especially very high levels of stunting by 24 months<sup>(12)</sup>. Prevalence of  
270 stunting in our study between HEU and HUU infants was similar to one found in a study of Ugandan  
271 infants enrolled in the PMTCT program<sup>(20)</sup>. Our bivariate analysis of the prevalence of stunting and  
272 underweight between HEU and HUU infants is therefore consistent with the larger body of literature.  
273 However, other studies conducted in sub-Saharan Africa did not find any differences in nutritional  
274 outcomes between HEU and HUU infants<sup>(19, 40-41)</sup>. It was found that HEU infants though born slightly  
275 smaller compared to HUU infants, were able to quickly catch up in weight and length<sup>(19, 41-42)</sup>. This  
276 lack of difference in growth patterns was attributed to higher levels of breastfeeding and/or effective  
277 counselling for feeding choices in HEU infants<sup>(19, 41)</sup>. In the current study HEU infants were more  
278 likely to be formula fed than breastfed compared to HUU infants. This may have contributed to their  
279 poor growth compared to HUU infants, since poor growth is linked to no or sub-optimal breastfeeding  
280<sup>(38, 43)</sup>. It is important to note that our regression modelling indicated that mode of feeding in the first  
281 year of life, was not a statistically significant predictor of undernutrition. However, these studies were  
282 conducted before ART was widely available to HIV positive women, therefore this may have resulted  
283 in no difference in growth between HEU and HUU infants<sup>(19, 41-42)</sup>. Other feeding practices such as  
284 age of introduction of complimentary feeding (weaning), average energy and protein intake and  
285 dietary diversity were not significantly different between HEU and HUU infants. Dietary diversity  
286 was poor in both groups of infants because majority of infants did not consume a variety of foods in  
287 the 24 hours preceding the study. Dietary diversity is an important indicator of the quality of the diet  
288 as opposed to the quantity of the food served<sup>(26, 33)</sup>.

289 HEU infants in this study were vulnerable to poorer nutritional outcomes, especially stunting because  
290 even after adjusting for other variables, HIV-exposure remained a strong predictor for stunting. This  
291 finding is consistent with results from a number of studies<sup>(18, 44-45)</sup>. A study, conducted in Tanzania  
292 found a lower length for age in HEU compared to HUU infants at three and six months<sup>(44)</sup>. A higher

293 risk of stunting in HEU compared to HUU infants has serious implications because stunting is  
294 associated with poorer psychomotor and mental development in HEU infants <sup>(45)</sup>. This may affect the  
295 future potential development of these infants, especially if stunting is not reversed within the first  
296 1000 days <sup>(46-48)</sup>. Factors such as exposure to ART during pregnancy, poor sanitation and infections  
297 in infants especially diarrhoea may account for the increased risk of stunting in HEU compared to  
298 HUU infants <sup>(26, 46)</sup>. In studies where poor growth was associated with HIV-exposure it was found  
299 that HEU infants had lower birthweight compared to HUU infants <sup>(14, 18, 40)</sup>. In the current study, HEU  
300 infants had lower birthweight compared to HUU infants, however this did not reach statistical  
301 significance. This is in contrast with a number of studies where HEU infants are more likely to be  
302 smaller at birth compared to HUU infants <sup>(49, 11, 44, 46)</sup>. Interestingly, low birthweight was a strong and  
303 consistent predictor for poor nutritional status (underweight, stunting and wasting). Infants with low  
304 birthweight tend to be more vulnerable to poor nutrition and/or diseases effect <sup>(14,18)</sup>. The findings of  
305 the current study show that birthweight is a more powerful predictor of later nutritional status than  
306 nutrient intakes from complementary feeds, breastmilk versus formula feeding, household and  
307 environmental factors including number of people living in a household, primary water source and  
308 income level. Even though birth length was not significantly lower in HEU compared to HUU infants,  
309 birth length remained a predictor for underweight, indicating that a lower birth length increased the  
310 risk of underweight in these infants. This is consistent with findings from some studies where birth  
311 length is a significant intermediary of growth in infants <sup>(44, 49-50)</sup>.

312 Consistent with a number of studies it was found that mother/care giver's education level was a  
313 predictor for stunting after adjustment for other variables <sup>(26, 12, 18, 51)</sup>. In addition, HIV positive  
314 mothers were significantly older than HIV negative mothers. Younger age and higher education level  
315 are associated with better nutritional outcomes because these caregivers tend to have more knowledge  
316 about optimal feeding, hygiene and child caring practices <sup>(12, 18, 51-52)</sup>. These caring practices may  
317 especially be relevant in settings where HEU infants tend to formula feed <sup>(18)</sup>. It was also found that  
318 HEU infants had significantly more siblings than HUU infants. A higher number of siblings is  
319 associated with poor nutritional outcomes in children <sup>(53)</sup>. Although growing up in a household where  
320 another child had died was a significant predictor of the risk of underweight in univariate analysis,  
321 after adjusting for potential confounding factors there was no relationship between the number of  
322 deceased siblings and risk of stunting, wasting or underweight.

323 Other determinants of nutritional outcomes in these infants included residing in Kweneng East district  
324 and Apgar score. Infants who resided in Kweneng East had a lower risk of wasting compared to those  
325 in other districts. It is should be noted that Kweneng East district was the only district where growth  
326 and health monitoring services were still offered in the main and primary hospital. Other districts

327 have moved these services to smaller clinics and/or health posts. Therefore, infants in Kweneng East  
328 district may have benefited from having close access to a multidisciplinary team of health  
329 professionals such as paediatricians and dietitians. These health-care workers are not typically  
330 accessible in smaller clinics. Accessibility to specialised care is highly relevant to wasting because  
331 wasting is an acute form of undernutrition, characterised by rapid weight loss due to acute inadequate  
332 intake and/or disease <sup>(54)</sup>. Therefore, infants in Kweneng East district were more likely to have  
333 accessed swift and specialised care upon being diagnosed with wasting compared to other districts.  
334 A higher Apgar score increased the risk of wasting in these infants almost two-fold. This was not  
335 expected because a higher Apgar score is associated with better nutritional outcomes <sup>(55)</sup>. However, a  
336 study in Asian Indian infants found that Apgar is a poor prognosis for growth and development in  
337 infants <sup>(56)</sup>.

338 It is important to note the following limitations about the current study. We have only considered the  
339 impact of HIV exposure, infant feeding, maternal and household factors upon nutritional status using  
340 the extreme outcome measures of stunting, wasting and underweight as determined by anthropometry  
341 and reference to WHO cut-offs for  $z$  scores. Indices such as micronutrient deficiencies were not  
342 included and we also did not focus on lower variance from cutoffs in terms of growth. Contribution  
343 of HIV-exposure may be greater at these subclinical levels and thus the  $z$  scores may be lower in HEU  
344 compared to HUU infants. Due to the cross-sectional study design, we did not have access to maternal  
345 nutrition and health indicators variables such as weight, height, CD4 count and use of ART pre-and  
346 post-natally. There is also a possibility, albeit a limited one, that some of the infants who were  
347 classified as HEU may have been HIV-infected after 6 weeks, since testing of HIV in these infants in  
348 Botswana is done at 6 weeks, post weaning if the mother was breastfeeding (6 months) and at 18  
349 months. Some of the infants in our study were not yet 18 months, at the time of data collection.  
350 However, a majority of these infants were formula feeding, therefore it was highly unlikely that they  
351 would have seroconverted. In addition, we have to acknowledge the cross-sectional nature of this  
352 study especially in regards to HIV-exposure and nutritional outcomes. Longitudinal studies are  
353 therefore required to elicit more data which will allow us to disentangle feeding modalities from HIV-  
354 exposure and also to derive more information on maternal nutrition and health during pregnancy.

355 PMTCT strategies in Botswana needs to be refined, so that optimal nutritional outcomes in HEU  
356 infants are prioritised in addition to prevention of MTCT of HIV. This can be achieved by integrating  
357 nutrition-specific and -sensitive interventions into this program. This will ensure equitable and  
358 optimal growth in HEU and HUU infants during their first 1000 days. Botswana as a country in terms  
359 of its health care system infrastructure, PMTCT strategies and growth surveillance for infants is in a  
360 good position to effect these significant changes, and thus improve population health.

361 In Botswana, HEU infants aged 6-24 months have poor nutritional status compared to HUU infants.  
362 Although mode of feeding was not a statistically significant factor determining risk of undernutrition,  
363 HEU infants tended to formula feed while HUU infants tended to breastfeed for the first twelve  
364 months of life. Therefore, HEU infants are missing out on the well documented benefits of  
365 breastfeeding. In order to increase breastfeeding levels in HEU infants there is need to review the  
366 current Botswana government's infant feeding policy in order to align with the new 2016  
367 recommendations by WHO. Furthermore, this study demonstrated that the strongest predictor of  
368 nutritional outcomes is birthweight, therefore strategies designed to optimise infants' nutritional  
369 status in the first 1000 days should aim to improve birthweight.

370

### 371 **Running title**

372 HIV-exposure and nutritional status in infants.

373

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## 395 **Transparency Declaration**

396 The lead author confirms that the manuscript is an honest, accurate and transparent account of the  
397 study being reported and that no important aspects of the study have been omitted and that any  
398 discrepancies from the study as planned have been explained. The reporting of this work is  
399 compliant with STROBE guidelines.

400

## 401 **References**

- 402           1. WHO, World Health Organisation (2015). Global Health Observatory data repository (1990-  
403           present). [http://www.who.int/gho/child\\_health/en/](http://www.who.int/gho/child_health/en/) (accessed 25<sup>th</sup> April 2016).
- 404           2. Black RE, Victoria CS, Walker SP et al. Maternal and child undernutrition and overweight  
405           in low-income and middle income countries. *Lancet*. 2013; 382: 427- 451.
- 406           3. Morris SS, Cogill B, Uauy R et al. Effective international action against undernutrition: why  
407           has it proven so difficult and what can be done to accelerate progress? *Lancet*. 2008; 371:  
408           608-621.
- 409           4. Victoria GS, Onis de M, Hallal PC et al. Worldwide timing of growth faltering: revisiting  
410           implications for interventions. *Paediatr*. 2010; 125: 473- 480.
- 411           5. WHO, World Health organisation (2016). Botswana Statistics Summary (2000- present).  
412           <http://apps.who.int/gho/data/node.country.country-BWA> (accessed 25<sup>th</sup> April 2016).
- 413           6. CSO, Central Statistics Office Botswana (2013). Botswana Aids Impact Survey IV (BAIS  
414           IV). Gaborone, Botswana.
- 415           7. MOH, Ministry of Health Botswana (2012). Botswana National HIV and AIDS Treatment  
416           Guidelines. Gaborone, Botswana.
- 417           8. UNAIDS (2016). *Global plan: on the fast track to an AIDS-free generation*. Geneva,  
418           Switzerland.
- 419           9. Rosala-Hallas A, Bartlett JW, Filteau S. Growth of HIV-exposed uninfected compared with  
420           HIV-unexposed, Zambian children: Longitudinal analysis from infancy to school age. *BMC*  
421           *Paedtr*. 2017; 17:80.
- 422           10. Filteau S. The HIV-exposed, uninfected African child. *Trop Med and Inter Health*. 2009; 14  
423           (3): 276-287.

- 424 11. McGrath CJ, Nduati R, Richardson BA et al. The prevalence of stunting is high in HIV-1  
425 exposed uninfected infants in Kenya. *J Nutr.* 2012; 142: 757 – 763.
- 426 12. Rollins NC, Ndirangu J, Bland RM et al. Exclusive breastfeeding, diarrhoeal morbidity and  
427 all-cause mortality in infants of HIV-infected and HIV uninfected mothers: an intervention  
428 cohort study in Kwazulu Natal, South Africa. *Plos ONE.* 2013; 8 (12): 1-10.
- 429 13. Shapiro RL, Lockman S, Kim S et al. Infant morbidity, mortality, and breast milk  
430 immunologic profiles among breast-feeding HIV –infected and HIV-Uninfected women in  
431 Botswana. *J Infec Dis.* 2007; 196:562 -569.
- 432 14. Marinda E, Humphrey JH, Iliff PJ et al. Child Mortality According to Maternal and Infant  
433 HIV Status in Zimbabwe. *Paed Infec Dis J.* 2007; 26: 519 -526.
- 434 15. Brahmabhatt H, Kigozi G, Wabwire-Mangen, F et al. Mortality in HIV-Infected and  
435 Uninfected Children of HIV-Infected and Uninfected Mothers in Rural Uganda. *J Acquir  
436 Immune Defic Syndr.* 2006; 41: 504-508.
- 437 16. Taha TE, Kumwenda NI, Broadhead RL et al. Mortality after the first year of life among  
438 human immunodeficiency virus type 1-infected and uninfected children. *Paed Infec Dis J.*  
439 1999; 18 (8): 689-694.
- 440 17. UNICEF, United Nations Children’s Fund (2013). Improving Child Nutrition: the  
441 achievable imperative for global progress. Report number: E.13.XX.4. New York:  
442 UNICEF.
- 443 18. McDonald CM, Kupka R, Manji KP et al. Predictors of stunting, wasting and underweight  
444 among Tanzanian children born to HIV-infected women. *HEUrop J Clin Nutr.* 2012;  
445 66:1265 -1276.
- 446 19. Patel D, Bland R, Coovadia H et al. Breastfeeding, HIV status and weights in South African  
447 Children; a comparison of HIV-exposed and unexposed children. *AIDS.* 2010; 24:437 -445.
- 448 20. Magezi SR, Kikafunda J, Whitehead R. Feeding and nutritional characteristics of infants on  
449 PMTCT programs. *J Trop Paedtr.* 2008; 55(1): 32-35.
- 450 21. Nnyepi M, Gobotswamang KSM, Codjia P. Comparison of estimates of malnutrition in  
451 children aged 0 -5 years between clinic based nutrition surveillance and national surveys. *J  
452 Public Health Policy.* 2011; 32: 281 – 282.
- 453 22. Chopra M and Rollins N. Infant feeding in time of HIV: rapid assessment of infant feeding  
454 policy and programmes in four African countries scaling up prevention of mother to child  
455 transmission programmes. *Arch Dis Child.* 2008; 93:288-291.
- 456

- 457 23. CSO, Central Statistics Office. (2015). Botswana. Population census atlas 2011. Gaborone,  
458 Botswana.
- 459
- 460 24. Botswana Tourism. (2013). Botswana Location.  
461 <http://www.botswanaturism.co.bw/location> . (accessed April 2017).
- 462 25. Green SB. How many subjects does it take to do a regression analysis? *Multi Behav*  
463 *Research*.1991; 26: 449-510.
- 464 26. Kimani-Murage EW, Norris SA, Pettifor JM, et al. Nutritional status and HIV in rural  
465 South African children. *BMC Paedtr*. 2011; 11: 1471 – 2431.
- 466 27. Akombi BJ, Agho AE, Merom D et al. Multilevel analysis of factors associated with  
467 wasting and underweight among children under-five years in Nigeria. *Nutrients*. 2017; 9  
468 (44): 1-17.
- 469 28. Mbwana HA, Kinabo J, Lambert C et al. Factors influencing stunting among children in  
470 rural Tanzania: an agro-climatic zone perspective. *Food Sec*. 2017; 1-17.
- 471 29. WHO-World Health Organisation (2009). *WHO child growth standards and the*  
472 *identification of severe acute malnutrition in infants and children*. Geneva: WHO
- 473 30. WHO, World Health Organisation (2005). WHO Anthro (version 3.2.2, January 2011)  
474 and macros. Geneva, Switzerland.
- 475 31. Raper N, Perloff B, Ingwersen L et al. An overview of USDA's dietary intake data system. *J*  
476 *Food Comp Analysis*. 2004; 17: 545-555.
- 477 32. Nightingale H, Walsh KJ, Oluput-Oluput P et al. Validation of triple pass 24-hour dietary  
478 recall in Ugandan children by simultaneous weighed food assessment. *BMC Nutr*. 2016; 2  
479 (56): 1-9.
- 480 33. WHO, World Health Organisation (2008). *Indicators for assessing infant and young child*  
481 *feeding*. Geneva, Switzerland.
- 482 34. Nutritics. Released 2011. Nutritics Professional Nutrition Analysis Software. Ireland.
- 483 35. SA FOODS (2015). South African Food Data System. [http://safoods-](http://safoods-apps.mrc.ac.za/foodcomposition/)  
484 [apps.mrc.ac.za/foodcomposition/](http://safoods-apps.mrc.ac.za/foodcomposition/) (accessed 24<sup>th</sup> February 2016).
- 485 36. Finglas PM, Roe MA, Pinchen HM et al. (2014) McCance and Widdowson's The  
486 Composition of Foods. Seventh summary edition. Cambridge: Royal Society of Chemistry.
- 487 37. IBM Corp. Released 2013. IBM SPSS Statistics for Windows. Version 22.0. Armonk, NY:  
488 IBM Corp.



- 489 38. Arpadi S, Fawzy A, Aldrovandi GM et al. Growth faltering due to breastfeeding cessation in  
490 uninfected children born to HIV-infected mothers in Zambia. *Am J Clin Nutr.* 2009; 90:344  
491 -353.
- 492 39. Bandara T, Hettiarachchi M, Liyanage C et al. current infant feeding practices and impact  
493 on growth in babies during the second half of infancy. *J Hum Diet.* 2015; 28:366-374.
- 494 40. Sherry B, Embree JE, Mei Z et al. Sociodemographic characteristics, care, feeding practices,  
495 and growth of cohorts of children born to HIV-1 seropositive and seronegative mothers in  
496 Nairobi, Kenya. *Tropic Med Inter Health.* 2000; 5 (10): 678-686.
- 497 41. Bailey R.C, Kamenga MC, Nsuami MJ et al. Growth of children according to maternal and  
498 child HIV, immunological and disease characteristics: a prospective cohort study in  
499 Kinshasa, Democratic Republic of Congo. *Inter Epidem Assoc.* 1999; 25: 532-540.
- 500 42. Isanaka S, Duggan C, Fawzi WW. Patterns of postnatal growth in HIV-infected and HIV  
501 exposed children. *Nutr Reviews.* 2009; 67 (6): 343-359.
- 502 43. Coovadia HM. & Bland RM. Preserving breastfeeding practice though the HIV pandemic.  
503 *Tropic Med and Inter Health.* 2007; 12: 1116-1133.
- 504 44. Wilkinson AL, Pedersen SH, Urassa M et al. Associations between gestational  
505 anthropometry, maternal HIV, and fetal and early infancy growth in a prospective  
506 rural/semi-rural Tanzanian cohort, 2012-13. *BMC Preg Childbirth.* 2015; 15 (277): 1-13.
- 507 45. McDonald CM, Manji KP, Kupka R et al. Stunting and wasting are associated with poorer  
508 psychomotor and mental development in HIV-exposed Tanzanian infants. *J Nutr.* 2013; 143:  
509 204-214.
- 510 46. Makasa M, Kasonka L, Chisenga M et al. Early growth of infants of HIV-infected and  
511 uninfected Zambian women. *Trop Med Inter Health.* 2007; 12 (5): 594-601.
- 512 47. Leroy JL, Ruel M, Habicht JP et al. linear growth deficit continues to accumulate beyond  
513 the first 1000 days in low- and middle-income countries: global evidence from 51 national  
514 surveys *Nutr.* 2014; 144: 1460-1466.
- 515 48. Victoria GC, de Onis M, Hallal PC et al. Worldwide timing of growth faltering: revisiting  
516 implications for interventions. *Paedtr.* 2010; 125 (3): 473-483.
- 517 49. Powis KM, Smeaton L, Hughes MD et al. In-utero triple antiretroviral exposure associated  
518 with decreased growth among HIV-exposed uninfected infants in Botswana. *AIDS.* 2016;  
519 30:211-220.
- 520 50. Sudfeld CR, Quanhong L, Chinyanga Y et al. Linear growth faltering among HIV-exposed  
521 uninfected children *AIDS.* 2017; 73 (2): 182-189.
- 522 51. Abuya BA, Ciera J. & Kimani-Murage E. Effect of mother's education on child's nutritional  
523 status in the slums of Nairobi. *BMC Paedtr* 2012; 12 (80): 1-10

- 524 52. Huynh D.T.T, Estorninos E, Capeding R.Z et al. Longitudinal growth and health outcomes  
525 in nutritionally at-risk children who received long-term nutritional intervention. *J Hum*  
526 *Nutr Diet.* 2015; 28:623-635.
- 527 53. Biswas S. & Bose K. Effect of number of rooms and sibs on nutritional status among rural  
528 Bengalee preschool children from eastern India. *Coll Antropol.* 2011; 35:1017-1022.
- 529 54. WHO, World Health Organisation (2017).  
530 [http://www.who.int/nutrition/topics/moderate\\_malnutrition/en/](http://www.who.int/nutrition/topics/moderate_malnutrition/en/) . (accessed April 2017)
- 531 55. Trivedi DJ, Shindhe V, Rockhade CJ. Influence of maternal nutrition during pregnancy on  
532 developmental outcome in first 30 days of independent neonatal life. *Inter J Clin Biochem*  
533 *Research.* 2016; 3(4): 371-375.
- 534 56. Lee CH, Ramachandran P, Madan A et al. Morbidity risk at birth fir Asian Indian small for  
535 gestational age infants. *Am J Public Health.*2010; 100 (5): 820-822.
- 536